



A radiomics model of contrast-enhanced computed tomography for predicting post-acute pancreatitis diabetes mellitus

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Background: Diabetes mellitus can occur after acute pancreatitis (AP), but the accurate quantitative methods to predict post-acute pancreatitis diabetes mellitus (PPDM-A) are lacking. This retrospective study aimed to establish a radiomics model based on contrast-enhanced computed tomography (CECT) for predicting PPDM-A.

Methods: A total of 374 patients with first-episode AP were retrospectively enrolled from two tertiary referral centers. There were 224 patients in the training cohort, 56 in the internal validation cohort, and 94 in the external validation cohort, and there were 86, 22, and 27 patients with PPDM-A in these cohorts, respectively. The clinical characteristics were collected from the hospital information system. A total of 2,398 radiomics features, including shape-based features, first-order histogram features, high order textural features, and transformed features, were extracted from the arterial- and venous-phase CECT images. Intraclass correlation coefficients were used to assess the intraobserver reliability and interobserver agreement. Random forest-based recursive feature elimination, collinearity analysis, and least absolute shrinkage and selection operator (LASSO) were used for selecting the final features. Three classification methods [eXtreme Gradient Boosting (XGBoost), Adaptive Boosting, and Decision Tree] were used to build three models and performances of the three models were compared. Each of the three classification methods were used to establish the clinical model, radiomics model, and combined model for predicting PPDM-A, resulting in a total of nine classifiers. The predictive performances of the models were evaluated by the area under the receiver operating characteristic curve (AUC), accuracy, sensitivity, specificity, positive predictive value, negative predictive value, and F1-score.

Results: Eleven radiomics features were selected after a reproducibility test and dimensionality reduction. Among the three classification methods, the XGBoost classifier showed better and more consistent performances. The AUC of the XGBoost's radiomics model to predict PPDM-A in the training, internal, and external cohorts was good (0.964, 0.901, and 0.857, respectively). The AUC of the XGBoost's combined model to predict PPDM-A in the training, internal, and external cohorts was good (0.980, 0.901, and 0.882,

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respectively). The AUC of the XGBoost's clinical model to predict PPDM-A in the training, internal, and external cohorts did not perform well (0.685, 0.733, and 0.619, respectively). In the external validation cohort, the AUC of the XGBoost's radiomics model was significantly higher than that of the clinical model (0.857 vs. 0.619, $P < 0.001$), but there was no significant difference between the combined and radiomics models (0.882 vs. 0.857, $P = 0.317$).

Conclusions: The radiomics model based on CECT performs well and can be used as an early quantitative method to predict the occurrence of PPDM-A.

Keywords: Radiomics; X-ray computed tomography (X-ray CT); acute pancreatitis (AP); diabetes mellitus

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Introduction

Acute pancreatitis (AP) is a common acute abdominal condition with an increasing incidence worldwide (1). Although AP is often considered a self-limiting disease, there is growing evidence that a new-onset disturbance in glucose metabolism can occur following AP (2). Recently, there has been growing interest in post-acute pancreatitis diabetes mellitus (PPDM-A). According to a 2014 systematic review and meta-analysis, 23% of AP patients develop new-onset diabetes during follow-up (3). A 2015 national population-based study showed that AP patients had greater than twofold higher risk of diabetes mellitus than individuals without AP (4). The prevalence of PPDM-A has almost tripled during the past 10 years, and its incidence is expected to increase by 2.8% per year, reaching 15.8 per 100,000 general population by 2050 (5,6). PPDM-A confers a higher risk of pancreatic cancer mortality and shorter life expectancy than type 2 diabetes (7-9). Therefore, it would be useful to develop an effective method for screening PPDM-A.

Previous studies have mainly examined the clinical risk factors for predicting PPDM-A (10-13). However, the simple clinical models showed low accuracy and stability and lacked individual specificity, which could be caused by the differences in research methods and clinical factors. Although imaging examinations performed well in diagnosing AP, there have been few imaging studies on the risk factors of PPDM-A. These studies have only analyzed the effects of intra-pancreatic fat deposition, Balthazar's score, and pancreatic necrosis on the risk of developing PPDM-A (14-16). It has found that excess intra-pancreatic

fat deposition is present in patients with PPDM-A, which can destroy pancreatic cells, impair pancreatic endocrine function, and ultimately increase the risk of pancreatic cancer (17). However, conventional imaging examinations are unable to analyze pancreatic microscopic changes at the molecular and cellular levels. Moreover, the existing predictive models only contain the traditional morphological features visible to the naked eye and do not include the potential subtle features that reflect disease heterogeneity. Therefore, accurate quantitative methods to predict PPDM-A are still lacking.

Quantitative imaging features are extracted from digital images in a high-throughput manner and changes in the human body at the tissue, cellular, and gene level can therefore be reflected by radiomics (18,19). Studies have found that radiomics models can be used to evaluate the severity of AP and indicate the prognosis and evolution of pancreatitis (20-22). One previous study developed a computed tomography (CT)-based radiomics nomogram to predict the incidence of PPDM-A, which was a single-center study with a small sample size, and only radiomic features were extracted from unenhanced plain CT images, ignoring potentially valuable information (23). Therefore, we established a radiomics model based on the arterial- and venous-phase contrast-enhanced CT (CECT) images at the first onset of AP from two tertiary care centers and compared it with a clinical model and a combined model based on the clinical and radiomic features to evaluate its capability to predict PPDM-A. We present this article in accordance with the TRIPOD reporting checklist (available at <https://qims.amegroups.com/article/view/10.21037/qims-23-1232/rc>).

Methods

Patients

This retrospective, bicentric study was conducted in accordance with the Declaration of Helsinki (as revised in 2013). The study was approved by the Institutional Review Board of the Second Affiliated Hospital of Chongqing Medical University and the First Affiliated Hospital of Chongqing Medical University (No. 315, dated December 30, 2021), and individual consent for this retrospective analysis was waived. We retrospectively recruited AP patients who had been hospitalized at institution 1 (the Second Affiliated Hospital of Chongqing Medical University) in the period from August 2014 to August 2022 and at institution 2 (the First Affiliated Hospital of Chongqing Medical University) from January 2018 to August 2022. The baseline information and imaging data were collected and analyzed.

Depending on the occurrence of diabetes after AP, the patients were divided into PPDM-A group and non-PPDM-A group (24,25). In PPDM-A group, the patients with a first onset of AP were followed up for the development of diabetes mellitus from August 2014 to August 2023. In non-PPDM-A group, the patients with a first AP onset were enrolled between August 2014 and August 2022 and were followed up through August 2023 via the analysis of medical records and telephone interviews to verify that diabetes mellitus did not occur during the follow-up period. In PPDM-A group, the mean interval time between the first onset of AP and the diagnosis of diabetes was 29 ± 24.6 months (3–120 months). In non-PPDM-A group, the mean follow-up time was 59.8 ± 28.1 months (12.3–134.8 months).

According to the 2012 Atlanta consensus criteria (26), AP was diagnosed based on the identification of at least two of the following three signs: (I) the presence of typical abdominal pain; (II) elevated serum amylase or lipase levels at least three times greater than the upper limit of normal; and (III) characteristic imaging findings of AP.

Diabetes was diagnosed by recording at least one of the following (27): (I) fasting plasma glucose levels ≥ 126 mg/dL (7.0 mmol/L); (II) plasma glucose ≥ 11.1 mmol/L 2 hours after a 75 g oral glucose load; (III) random blood glucose ≥ 200 mg/dL (11.1 mmol/L) accompanied by diabetes symptoms; and (IV) glycosylated haemoglobin $\geq 6.5\%$ (48 mmol/dL).

The inclusion and exclusion criteria for PPDM-A group were as follows: inclusion criteria: (I) patients were

diagnosed with diabetes mellitus >90 days after the date of first pancreatitis diagnosis (24,25,28); (II) upper abdominal CECT examinations performed within 7 days of the onset of AP; and (III) inpatients. Exclusion criteria: (I) having any history of diabetes; (II) elevated blood glucose during hospitalization; (III) high blood glucose within 90 days after discharge; (IV) acute onset of chronic pancreatitis; (V) cancer or severe chronic wasting diseases; (VI) loss to follow-up; (VII) unsatisfactory images or incomplete medical records; and (VIII) less than 18 years of age.

The inclusion criteria applied to the non-PPDM-A were based on the diagnosis of AP and (I) a detailed admission index at the first AP onset and (II) there was no abnormal blood sugar index, after a detailed review of medical records or follow-up telephone consultations through August 2023. The exclusion criteria were as follows: (I) acute onset of chronic pancreatitis; (II) cancer or severe chronic wasting diseases; (III) complications resulting from abnormal pancreatic anatomical features; (IV) loss to follow-up; and (V) unsatisfactory images or incomplete medical records.

A total of 374 patients with first onset AP were recruited (mean age, 48 ± 15 years; age range, 18–86 years; 231 men), comprising 239 participants in the non-PPDM-A group and 135 participants in the PPDM-A group. The patients from institution 1 were randomly assigned to the training cohort and the internal validation cohort at a ratio of 8:2, while those from institution 2 served as the external validation cohort. Twelve clinical characteristics, namely age, gender, disease severity based on the 2012 Revised Atlanta Criteria (26), CT severity index (CTSI), extra-pancreatic inflammation on CT (EPIC) (29), etiology, hospital stay, pancreatic necrosis, the number of recurrences, follow-up time, and smoking and drinking status, were collected from the hospital information system. The classification of etiological factors and the definition of recurrent AP are shown in Section S1 in [Appendix 1](#), and the study flowchart is presented in [Figure 1](#).

CECT image collection

All of the patients underwent an abdominal CECT examination within 7 days after the onset of AP symptoms with one of three multidetector row CT systems. We retrieved anonymous arterial- and venous-phase CECT images from the picture archiving and communication system for feature extraction. The detailed image collection information is provided in Section S2 in [Appendix 1](#).

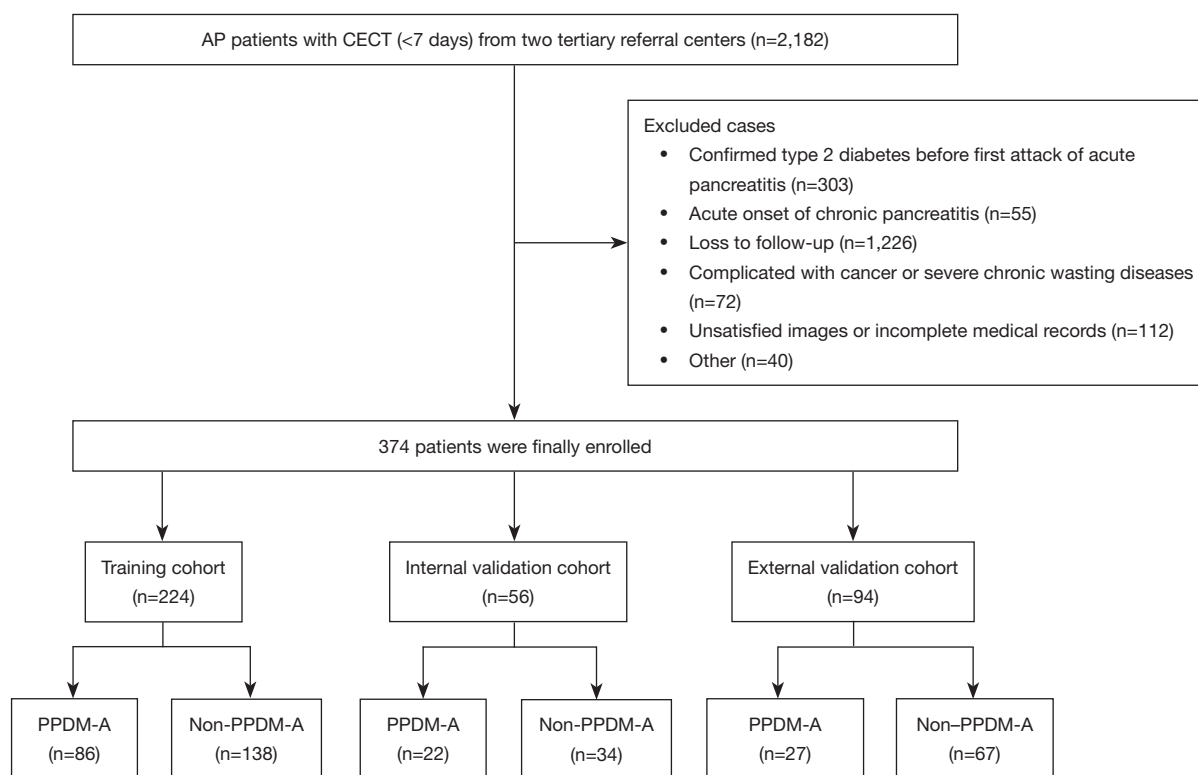


Figure 1 Flowchart of patient recruitment in this study. AP, acute pancreatitis; CECT, contrast-enhanced computed tomography; PPDM-A, post-acute pancreatitis diabetes mellitus.

Image segmentation, preprocessing, and feature extraction

Two abdominal radiologists with 8 and 12 years of experience used the IntelliSpace Discovery platform (ISD, Philips Healthcare, Best, the Netherlands) independently, without knowing the patient's clinical outcomes. On the ISD platform, a three-dimensional region of interest (3DROI) containing the whole pancreas was semiautomatically generated by delineating the pancreatic boundaries on the CECT image, including pancreatic necrosis but avoiding the bile ducts and blood vessels (30). Then, the plugin-radiomics 2.1 on the ISD platform was used to extract radiomics features for each 3DROI. A total of 1,199 radiomics features were extracted, as shown in Section S3 in [Appendix 1](#) and [Figure S1](#). For each patient, we used two different 3DROIs from CECT images of the arterial and portal venous phases and integrated all 2,398 radiomics features from the two 3DROIs. To ensure the repeatability of the results, the Min–Max normalization and resampling were used to preprocess the data and images, respectively (Section S4 in [Appendix 1](#)). The detailed radiomics process is presented in [Figure 2](#).

Intraobserver reliability and interobserver agreement

Two abdominal radiologists selected CECT images of random 30 patients, without knowing the patient's clinical outcomes, to assess intraobserver and interobserver agreement for feature extraction. To evaluate intraobserver reliability, observer 1 used the same method to delineate 3DROI and extract radiomics features twice within 7 days. To evaluate interobserver agreement, observer 2 delineated the 3DROI only once, and we compared the obtained features with the features extracted after the first delineation by observer 1. Intraclass correlation coefficients (ICCs) were used to assess the intraobserver reliability and interobserver agreement. The features with an ICC value of above 0.75 were considered stable, while the remaining features were considered unstable and eliminated.

Dimensionality reduction and feature selection

To avoid the dimensionality disarray and reduce the bias of radiomics features, the following measures were taken to select features in the training group. First, random forest-

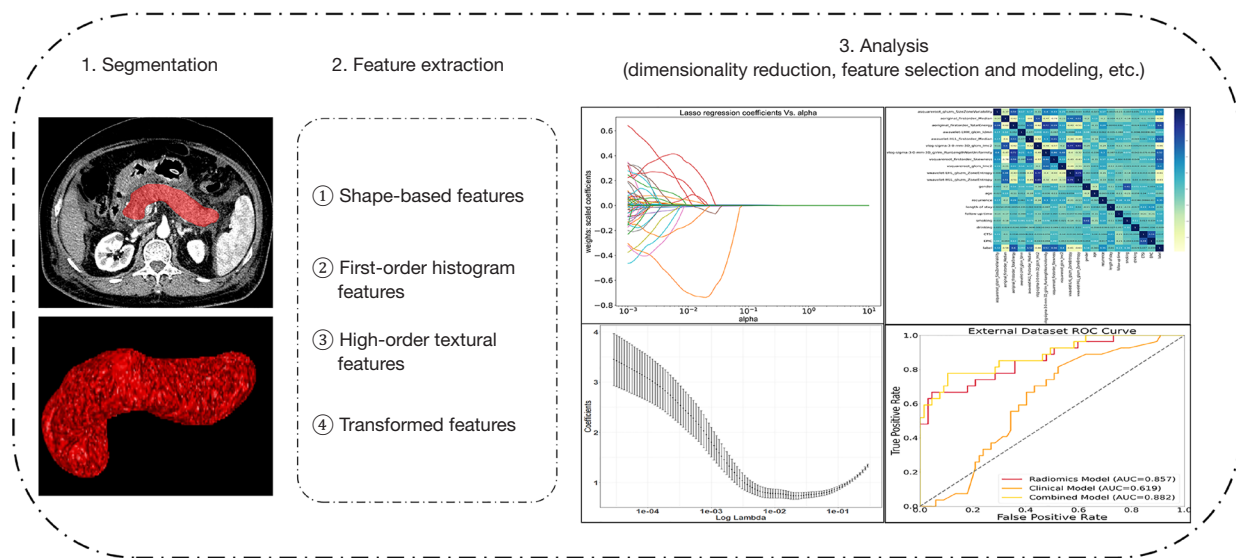


Figure 2 Workflow of radiomics process in this study. CTSI, computed tomography severity index; EPIC, extra-pancreatic inflammation on computed tomography; ROC, receiver operating characteristic curve; AUC, area under the ROC curve.

based recursive feature elimination was used to select features. Then, collinearity analysis was performed on the features selected by the random forest algorithm. The features with Pearson correlation coefficient (r) greater than 0.9 were randomly deleted to remove redundancy. Finally, the least absolute shrinkage and selection operator (LASSO) was used for dimensionality reduction to select the optimal features by performing regularization and variable selection so as to improve the predictive accuracy and reliability of the radiomics model. The idea was to control the amounts of features by introducing L1 regularization terms and to reduce the coefficients of some unimportant features. By adjusting the regularization parameter λ , the weight of regularization terms can be controlled. And the best regression function of the model can be selected while filtered out the sparse features.

Modeling

Three classification methods [eXtreme Gradient Boosting (XGBoost), Adaptive Boosting (AdaBoost), and Decision Tree] were built in three models to compare the performances. Fivefold cross-validation was applied to tune hyperparameters for better model performance. The predictive performances of the models were compared based on the area under the receiver operating characteristic (ROC) curve (AUC), accuracy, sensitivity, specificity,

positive predictive value, negative predictive value, and F1-score. Each of the three models was modeled with three algorithms, resulting in a total of nine classifiers. The performances were compared both within and between the models.

Statistical analysis

Statistical analyses were performed using SPSS for Windows (version 22.0, Chicago, IL, USA). The data in different groups were tested for normality using the Shapiro-Wilk test. Continuous variables that approximated a normal distribution were expressed as mean \pm standard deviation; continuous variables that were skewed were expressed as median (interquartile range). Independent-sample t -test or Mann-Whitney U test was applied to compare continuous variables. Chi-squared test or Fisher's exact test was used to compare categorical variables. Radiomics feature analysis, modeling and visual analysis were performed using Python 3.9.13 and R (version 4.3.1, <https://www.r-project.org/>). Pandas 1.4.4 and Numpy 1.21.5 were used for data preprocessing and normalization. Scikit-learn 1.0.2 was used for feature selection methods, establishment of machine learning models, and drawing the ROC curve. R was used for visual analysis and basic data analysis since it has a powerful graphing function. DeLong test was used to compare the AUCs of these models. The P

values were two-sided, and P values lower than 0.05 were considered statistically significant.

Results

Patient characteristics

The detailed clinical characteristics of the AP patients in the training, internal validation, and external validation cohorts are shown in *Table 1*. The prevalence of PPDM-A was 38.39%, 39.29%, and 28.72% in the training, internal validation, and external validation cohorts, respectively, with no statistically significant difference among the three cohorts ($P=0.236$). Of the 12 clinical characteristics measured, only the follow-up time and the number of recurrences significantly differed between the PPDM-A group and the non-PPDM-A group for the training cohort ($P=0.018$ and $P<0.001$, respectively) and the internal validation cohort ($P<0.001$ and $P=0.004$, respectively). The gender ($P=0.017$) and smoking status ($P=0.042$) significantly differed between the PPDM-A group and the non-PPDM-A group only in the training cohort. All of the clinical characteristics were normalized to values between 0 and 1 using Min–Max normalization, and nine features were selected by LASSO. Three clinical characteristics, namely, etiology, disease severity, and pancreatic necrosis, were excluded.

Intraobserver reliability and interobserver agreement

Regarding intraobserver reliability, a total of 2,130 features reached satisfactory agreement, with a mean ICC value of 0.957 and a range of 0.752–1.000, as shown in *Figure 3A*. The remaining 268 features were eliminated. For interobserver agreement, 1,807 features reached satisfactory agreement, with a mean ICC value of 0.946 and a range of 0.75–1.000, as shown in *Figure 3B*. The 591 features with unsatisfactory agreement were eliminated. Among the eliminated features, 140 features were identical between the intraobserver and interobserver agreement. Finally, 719 features were excluded, and the remaining 1,679 features were kept for further analyses.

Dimensionality reduction and feature selection

First, 505 radiomics features were selected by random forest-based recursive feature elimination. Next, 193 radiomics features remained after checking the collinearity. Then, 11 radiomics features were selected by LASSO

regression, and the value of the best-tuned regularization parameter (λ) was 0.01. Finally, eleven radiomics features were selected (Section S5 in *Appendix 1* for detailed feature information).

Modeling

Overall, among the three classification methods, the XGBoost classifier showed better and more consistent performances. In the external validation cohort, the XGBoost method achieved an AUC of 0.857 [95% confidence interval (CI): 0.855–0.981] in the radiomics model and an AUC of 0.882 (95% CI: 0.844–0.981) in the combined model. The AdaBoost method achieved AUC values of 0.837 (95% CI: 0.870–0.985) and 0.788 (95% CI: 0.852–0.981) in the radiomics and combined models, respectively. The Decision Tree method showed some AUC values of 0.5, meaning that the corresponding classification model was worthless. The detailed information on the AdaBoost and Decision Tree methods is shown in *Tables S1,S2*. Therefore, we further focused only on the best method's (XGBoost) performance among the three models.

The XGBoost's radiomics model performed well in predicting PPDM-A (*Table 2*). In the training cohort, the AUC and accuracy were 0.964 (95% CI: 0.896–0.967) and 0.964, respectively. In the internal validation cohort, the AUC and accuracy were 0.901 (95% CI: 0.849–0.964) and 0.875, respectively. In the external validation cohort, the AUC and accuracy were 0.857 (95% CI: 0.855–0.981) and 0.968, respectively.

The XGBoost's clinical model did not perform well in predicting PPDM-A (*Table 2*). In the training cohort, the AUC and accuracy were 0.685 (95% CI: 0.613–0.719) and 0.701, respectively. In the internal validation cohort, the AUC and accuracy were 0.733 (95% CI: 0.5–0.5) and 0.696, respectively. In the external validation cohort, the AUC and accuracy were 0.619 (95% CI: 0.50–0.64) and 0.596, respectively.

The XGBoost's combined model also showed good performance (*Table 2*). In the training cohort, the AUC and accuracy were 0.980 (95% CI: 0.916–0.982) and 0.978, respectively. In the internal validation cohort, the AUC and accuracy were 0.901 (95% CI: 0.789–0.963) and 0.893, respectively. In the external validation cohort, the AUC and accuracy were 0.882 (95% CI: 0.844–0.981) and 0.840, respectively.

By comparing the AUCs among the three XGBoost's models, we found that the radiomics model was significantly

Table 1 Clinical characteristics of the training cohort, internal validation cohort and external validation cohort

Characteristics	Institution 1						Institution 2		
	Training cohort (n=224)		P	Internal validation cohort (n=56)		P	External validation cohort (n=94)		
	PPDM-A (n=86)	Non-PPDM-A (n=138)		PPDM-A (n=22)	Non-PPDM-A (n=34)		PPDM-A (n=27)	Non-PPDM-A (n=67)	P
Gender			0.017*			0.757			0.074
Male	60 (69.77)	74 (53.62)		14 (63.64)	23 (67.65)		21 (77.78)	39 (58.21)	
Female	26 (30.23)	64 (46.38)		8 (36.36)	11 (32.35)		6 (22.22)	28 (41.79)	
Age (years)	46.97±15.31	48.88±15.18	0.362	50.32±13.54	49.03±15.04	0.746	51.37±15.88	46.15±16.03	0.16
Etiology			0.058			0.089			0.111
Biliary	21 (24.42)	56 (40.58)		4 (18.18)	15 (44.12)		11 (40.74)	33 (49.25)	
Hypertriglyceridemia	29 (33.72)	29 (21.01)		7 (31.82)	11 (32.35)		10 (37.04)	18 (26.87)	
Alcoholic	6 (6.98)	9 (6.52)		1 (4.55)	2 (5.88)		2 (7.41)	0	
Idiopathic	30 (34.88)	44 (31.88)		10 (45.45)	6 (17.65)		4 (14.81)	16 (23.88)	
CTSI	4 [3–6]	4 [3–4]	0.992	3.5 [3–4]	4 [2.75–4]	0.889	4 [3–6]	4 [3–6]	0.403
Disease severity			0.868			0.516			0.325
Mild	18 (20.93)	26 (18.84)		4 (18.18)	10 (29.41)		6 (22.22)	11 (16.42)	
Moderate	63 (73.26)	102 (73.91)		18 (81.82)	23 (67.65)		20 (74.07)	46 (68.65)	
Severe	5 (5.81)	10 (7.25)		0	1 (2.94)		1 (3.8)	10 (14.93)	
EPIC	5 [2–7]	5 [2–7]	0.447	4.5 [2.75–6]	4.5 [1–7]	0.858	5 [2–6]	5 [3–7]	0.449
Hospital stay (days)	13 [8–20]	15 [9–21]	0.243	13 [8.75–16.25]	11 [7–17.25]	0.608	12 [9–22]	18 [11–30]	0.083
Follow-up time (months)	22 [9–45]	15 [6–29]	0.018*	24.5 [18.5–37.5]	5.5 [3–20]	<0.001*	20 [8–36]	11 [5–27]	0.071
Pancreatic necrosis			0.428			0.613			0.595
Yes	24 (27.91)	32 (23.19)		4 (18.18)	6 (17.65)		8 (29.63)	20 (29.85)	
No	62 (72.09)	106 (76.81)		18 (81.82)	28 (82.35)		19 (70.37)	47 (70.15)	
Number of recurrences			<0.001*			0.004*			0.266
0 recurrence	23 (26.74)	91 (65.94)		5 (22.73)	22 (64.71)		11 (40.75)	40 (59.7)	
1 recurrence	41 (47.67)	35 (25.36)		6 (27.27)	8 (23.53)		14 (51.85)	22 (32.84)	
2 recurrences	12 (13.95)	6 (4.35)		6 (27.27)	3 (8.82)		1 (3.7)	3 (4.48)	
≥3 recurrences	10 (11.63)	6 (4.35)		5 (22.73)	1 (2.94)		1 (3.7)	2 (2.99)	
Smoking			0.042*			0.752			0.02*
Yes	39 (45.35)	44 (31.88)		10 (45.45)	14 (41.18)		15 (55.56)	20 (29.85)	
No	47 (54.65)	94 (68.12)		12 (54.55)	20 (58.82)		12 (44.44)	47 (70.15)	
Drinking			0.296			0.004*			0.542
Yes	26 (30.23)	33 (23.91)		7 (31.82)	24 (70.59)		6 (22.22)	19 (28.36)	
No	60 (69.77)	105 (76.09)		15 (68.18)	10 (29.41)		21 (77.78)	48 (71.64)	

The age is approximate a normal distribution which are expressed as mean ± standard deviation; CTSI, EPIC, hospital stay and follow-up time are skewed distribution which are expressed as median [interquartile range]. The gender, etiology, disease severity, pancreatic necrosis, number of recurrences, smoking and drinking status are expressed as number (frequency). *, P<0.05. PPDM-A, post-acute pancreatitis diabetes mellitus; CTSI, computed tomography severity index; EPIC, extra-pancreatic inflammation on computed tomography.

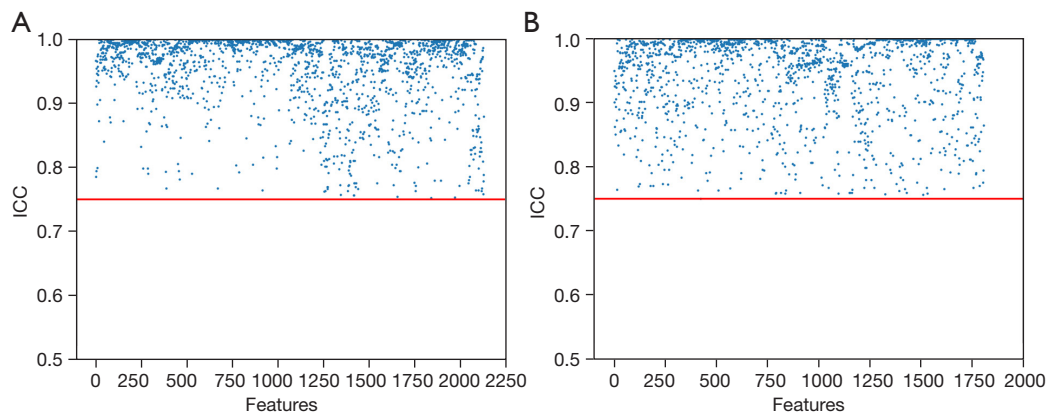


Figure 3 Evaluation of feature stability and intraobserver and interobserver agreement based on ICC. (A) All features presented good intraobserver agreement with ICC values higher than 0.75 (above the red cutoff line). (B) All features presented good interobserver agreement with ICC values higher than 0.75. ICC, intraclass correlation coefficient.

Table 2 The performance of the radiomics model, clinical model, and combined model in the training cohort, the internal validation cohort and the external validation cohort with the XGBoost method

Model	Training cohort			Internal validation cohort			External validation cohort		
	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined
AUC	0.685	0.964	0.980	0.733	0.901	0.901	0.619	0.857	0.882
(95% CI)	(0.613, 0.719)	(0.896, 0.967)	(0.916, 0.982)	(0.5, 0.5)	(0.849, 0.964)	(0.789, 0.963)	(0.50, 0.64)	(0.855, 0.981)	(0.844, 0.981)
Accuracy	0.701	0.964	0.978	0.696	0.875	0.893	0.596	0.968	0.840
Sensitivity	0.616 (53/86)	0.965 (83/86)	0.988 (85/86)	0.773 (17/22)	0.773 (17/22)	0.955 (21/22)	0.407 (11/27)	0.963 (26/27)	0.778 (21/27)
Specificity	0.754 (104/138)	0.964 (133/138)	0.971 (134/138)	0.647 (22/34)	0.941 (32/34)	0.853 (29/34)	0.672 (45/67)	0.970 (65/67)	0.866 (58/67)
PPV	0.609 (53/87)	0.943 (83/88)	0.955 (85/89)	0.586 (17/29)	0.895 (17/19)	0.808 (21/26)	0.333 (11/33)	0.929 (26/28)	0.700 (21/30)
NPV	0.759 (104/137)	0.978 (133/136)	0.993 (134/135)	0.815 (22/27)	0.865 (32/37)	0.967 (29/30)	0.738 (45/61)	0.985 (65/66)	0.906 (58/64)
F1-score	0.613	0.953	0.971	0.667	0.844	0.875	0.367	0.945	0.737

The brackets in the table show the numerator and denominator of the sensitivity, specificity, PPV and NPV. AUC, area under the receiver operating characteristic curve; CI, confidence interval; PPV, positive predictive value; NPV, negative predictive value.

better than the clinical model in the external validation cohort (0.857 *vs.* 0.619, $P < 0.001$), but there was no significant difference between the combined and radiomics models (0.882 *vs.* 0.857, $P = 0.317$). The ROC evaluation of the models is shown in *Figure 4*.

Discussion

In this study, a quantitative radiomics model based on CECT images was developed and validated, which provided a noninvasive and individualized method to predict PPDM-A. The radiomics model showed good performance in the external validation cohort, with an accuracy of 0.968,

suggesting that the clinical application of radiomics is promising in predicting PPDM-A.

The risk of PPDM-A increases with increasing follow-up time, the number of recurrences, and other clinical factors (31-33). In this study, we found significant differences in the follow-up time and the number of recurrences between the PPDM-A and non-PPDM-A groups in the training and internal validation cohorts. However, in the external validation cohort, the follow-up time was longer in the PPDM-A group than in the non-PPDM-A group, but the number of recurrences was not significantly different. In addition, the gender and smoking status were significantly different only in the training cohort. We speculate that

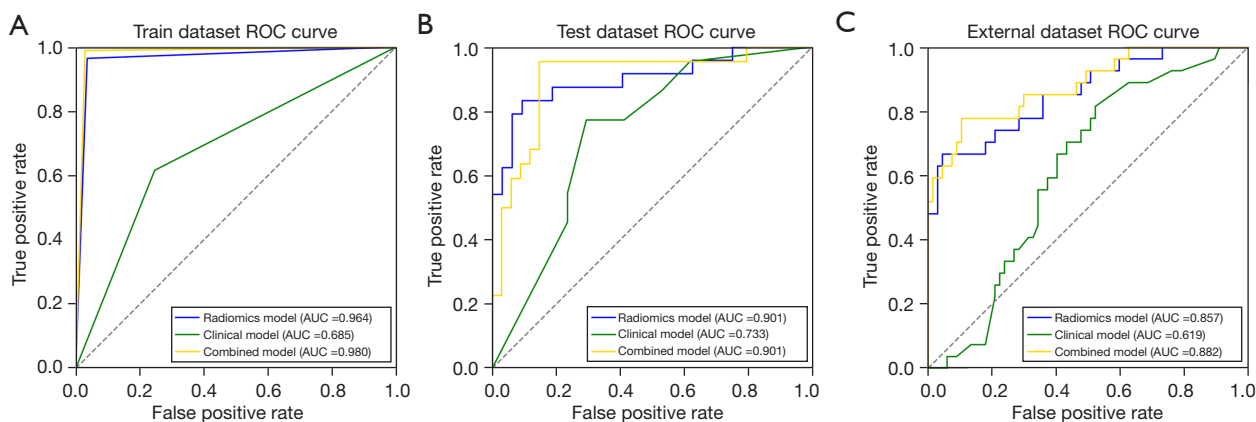


Figure 4 ROC of the radiomics model, clinical model, and combined model for predicting post-acute pancreatitis diabetes mellitus with the XGBoost's method. (A) ROC evaluation in training cohort. (B) ROC evaluation in internal validation cohort. (C) ROC evaluation in external validation cohort. ROC, receiver operating characteristic curve; AUC, the area under the ROC curve.

the reason for these results may be the small sample size. According to large population-based studies, the severity of the initial AP episode, pancreatic necrosis, and etiology are not related to the risk of PPDM-A (3,4,32), which is consistent with our findings. However, several studies have reported that pancreatic necrosis is an independent risk factor for PPDM-A (16,34). In clinical practice, PPDM-A can occur without pancreatic necrosis, suggesting that pancreatic necrosis is not the only reason for developing PPDM-A. Therefore, LASSO regression was used to exclude etiology, pancreatic necrosis, and severity so as to improve the accuracy of the clinical model. However, the accuracy of the clinical model to predict PPDM-A was only approximately 0.7, which was far from satisfactory.

In contrast, our proposed radiomics model showed obvious advantages in predicting PPDM-A compared with the clinical model. In this study, we analyzed CECT radiomics in patients with AP, and the results indicated that some patients were already at high risk of PPDM-A during the first episode of AP. There is evidence that stress hyperglycemia occurs in up to 70% of patients during the course of AP, and diabetes mellitus occurs in nearly 25% of patients after AP, suggesting that some patients experience a temporary or permanent impairment of pancreatic endocrine function during AP attacks (35,36). In recent years, it has been suggested that the pathogenesis of PPDM-A involves multiple factors. Previous studies have confirmed that chronic low-grade inflammation caused by elevated levels of proinflammatory cytokines during AP attacks is the key pathway of PPDM-A (37). In addition, proinflammatory cytokines-stimulated lipolysis plays an

important role in the pathogenesis of PPDM-A (38,39). It has been shown that intra-pancreatic fat deposition occurs in patients with PPDM-A, which may damage the Langerhans islet cells and impair insulin secretion (14,15). The dysfunction of the pancreas-gut-brain axis may also lead to PPDM-A (40,41). Both the enteroendocrine and brain cells can secrete the glucoregulatory peptides that affect the pancreatic islet cells, and the glucoregulatory peptides can affect the interaction between the nervous system and the gastrointestinal tract and regulate glucose metabolism (40). Therefore, PPDM-A may be caused by the changes in various factors in the microenvironment *in vivo*, and radiomics can reflect these changes to predict PPDM-A.

Zhong *et al.* (23) developed a CT-based radiomics nomogram for predicting PPDM-A incidence. The paper's authors conducted a single-center study with a small sample size and extracted radiomics features from only the unenhanced plain CT images, so potentially valuable information had been overlooked. And the AUC and accuracy of the CT-based radiomics nomogram to predict the PPDM-A incidence was only 0.743 and 0.815, respectively. Whereas our study collected arterial- and venous-phase CECT images from two tertiary referral centers, and provided an independent external validation cohort, providing more abundant image information and increasing the advantages of our results for clinical application. Besides, to reduce the influence of variable CT parameters on the stability of the radiomics features, we used resampling to preprocess the images. It has been shown that resampling can maintain feature stability by

optimizing gray-level discretization (42,43).

In our study, the radiomics model in the external validation cohort had a higher accuracy not only than the internal validation cohort but also the training cohort. We speculate that the reason for this situation may be due to the class imbalance. The prevalence of PPDM-A was higher in both the training and internal validation group (38.39% and 39.29%, respectively) than in the external validation group (28.72%). Besides, only the accuracy of the radiomics model was slightly higher. Therefore, we thought this was a normal phenomenon, which also can be found in the studies by the Du *et al.* and Ouyang *et al.* (44,45).

Our study has several limitations. First, the result might be slightly affected due to class imbalance. The imbalance learning techniques, such as the synthetic minority oversampling technique, undersampling, oversampling, are prone to overfitting problems and the generated sample may introduce noises (46). Although we did not adopt methods to deal with the class imbalance, the AUC values of our model were good eventually. In addition, many studies also did not take the approach of dealing with class imbalances and got satisfactory results (47,48). Second, although the follow-up period was relatively long, it is possible that PPDM-A did not occur within the follow-up period in some patients, which could have led to false-negative predictions. Besides, the clinical characteristics of AP patients were limited, which may have affected the prediction effect of PPDM-A in the comprehensive evaluation. Finally, the sample size was not large enough, which may have led to a slight selection bias. In the future, we will conduct a large prospective study to address these issues.

Conclusions

In conclusion, we proposed a radiomics model based on CECT images, providing a convenient and noninvasive method for predicting PPDM-A.

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Footnote

Reporting Checklist: The authors have completed the TRIPOD reporting checklist. Available at <https://qims.amegroups.com/article/view/10.21037/qims-23-1232/rc>

Conflicts of Interest: All authors have completed the ICMJE uniform disclosure form (available at <https://qims.amegroups.com/article/view/10.21037/qims-23-1232/coif>). R.H. reports that this research was funded by the Chengdu University of Traditional Chinese Medicine “Xinglin Scholars” Discipline Talents Research Promotion Plan (grant No. YYZX2021059). D.Z. reports that this research was sponsored by Natural Science Foundation of Chongqing, China (grant No. CSTB2023NSCQ-MSX0154) and the First Affiliated Hospital of Chongqing Medical University (grant No. CYYY-BSHPYXM-202305). The other authors have no conflicts of interest to declare.

Ethical Statement: The authors are accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved. The study was conducted in accordance with the Declaration of Helsinki (as revised in 2013). The study was approved by the institutional review board of the Second Affiliated Hospital of Chongqing Medical University and the First Affiliated Hospital of Chongqing Medical University (No. 315. Dated December 30, 2021), and individual consent for this retrospective analysis was waived.

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Appendix 1

S1 The classification of etiological factors and the definition of recurrent acute pancreatitis (AP)

S1.1 Classification of etiological factors

The etiology was divided into the following groups: (I) biliary, when gallstones were detected in the gallbladder or bile ducts in any image; (II) alcoholic, where 60 g or more alcohol was consumed per day for 5 years; (III) idiopathic, when no cause was found after a detailed examination; (IV) hypertriglyceridemia, when the level of fasting triglycerides was >1,000 mg/dL (11.3 mmol/L).

S1.2 Definition recurrent AP

Recurrent AP was defined as one or more recurrent AP episodes >3 months after the first AP in complete remission.

S2 Computed tomography (CT) image acquisition

All patients underwent contrast-enhanced computed tomography (CECT) upon abdominal imaging using one of the three following multidetector row CT systems: Aquilion ONE (Toshiba, Tokyo, Japan), Ingenuity CT (Philips Medical System), and Somatom Definition Flash (Siemens Healthineers). For the first two CT scanners, the main acquisition parameters were as follows: tube voltage of 120 kV (both), tube current of 250 mA (both), field of view (FOV) of 40×40 cm (both), matrix of 512×512 (both), collimation of 320×0.5 mm and auto, pitch of 0.87 and 1.015, and slice thickness of 5.0 mm (both). After a routine nonenhanced scan, arterial- and portal venous-phase CECT scans were performed after 25–30 and 48–50 s of delay following the intravenous administration of iodinated contrast material (Ultravist 370, Bayer Schering Pharma) at 1.5 mL/kg at a rate of 3 mL/s using a pump injector.

For Somatom Definition Flash, the main acquisition parameters were as follows: tube voltage of 100 kV, tube current of 318 mA, FOV of 33×33 cm, collimation of 128×0.6 mm, pitch of 0.8, and slice thickness of 5.0 mm. An automatic exposure control system (Care Dose 4D; Siemens Medical Solutions) was used when performing scanning. After a routine nonenhanced scan, arterial- and portal venous-phase CECT scans were performed after 25 and 40 s of delay following the intravenous administration of iodinated contrast material (Ultravist 370, Bayer Schering Pharma) at 1.5 mL/kg at a rate of 3.5–5 mL/s using a pump injector.

S3 Feature extraction

The details of extracted radiomics features are shown in *Figure S1*.

S.3.1 First-order statistics (N=19)

- ❖ Interquartile range;
- ❖ Skewness;
- ❖ Uniformity;
- ❖ Median;
- ❖ Energy;
- ❖ Robust mean absolute deviation;
- ❖ Mean absolute deviation;
- ❖ Total energy;
- ❖ Maximum;
- ❖ Root mean squared;
- ❖ 90th percentile;
- ❖ Minimum;
- ❖ Entropy;
- ❖ Standard deviation;

- ❖ Range;
- ❖ Variance;
- ❖ 10th percentile;
- ❖ Kurtosis;
- ❖ Mean.

S3.2 High-order statistics (N=60)

Gray level cooccurrence matrix (N=28)

- ❖ Sum variance;
- ❖ Homogeneity 1;
- ❖ Homogeneity 2;
- ❖ Cluster shade;
- ❖ Maximum probability;
- ❖ Idmn;
- ❖ Sum variance 2;
- ❖ Contrast;
- ❖ Difference entropy;
- ❖ Inverse variance;
- ❖ Entropy;
- ❖ Dissimilarity;
- ❖ Difference variance;
- ❖ Idn;
- ❖ Idm;
- ❖ Correlation;
- ❖ Autocorrelation;
- ❖ Sum entropy;
- ❖ Average intensity;
- ❖ Energy;
- ❖ Sum squares;
- ❖ Cluster prominence;
- ❖ Sum average;
- ❖ Imc2;
- ❖ Imc1;
- ❖ Difference average;
- ❖ Id;
- ❖ Cluster tendency.

Gray level run length matrix (N=16)

- ❖ Short run low gray level emphasis;
- ❖ Gray level variance;
- ❖ Low gray level run emphasis;
- ❖ Gray level non-uniformity normalized;
- ❖ Run variance;
- ❖ Gray level non-uniformity;
- ❖ Long run emphasis;
- ❖ Short run high gray level emphasis;
- ❖ Run length non-uniformity;
- ❖ Short run emphasis;
- ❖ Long run high gray level emphasis;

- ❖ Run percentage;
- ❖ Long run low gray level emphasis;
- ❖ Run entropy;
- ❖ High gray level run emphasis;
- ❖ Run length non-uniformity normalized.

Gray level size zone matrix (N=16)

- ❖ Gray level variance;
- ❖ Low intensity large area emphasis;
- ❖ High intensity small area emphasis;
- ❖ Small area emphasis;
- ❖ Large area emphasis;
- ❖ Zone variance;
- ❖ Size zone variability normalized;
- ❖ Low intensity small area emphasis;
- ❖ High intensity emphasis;
- ❖ Intensity variability normalized;
- ❖ Zone percentage;
- ❖ Low intensity emphasis;
- ❖ Size zone variability;
- ❖ Intensity variability;
- ❖ Zone entropy;
- ❖ High intensity large area emphasis.

S3.3 Shape-based features (N=14)

- ❖ Maximum 3D diameter;
- ❖ Compactness 2;
- ❖ Maximum 2D diameter slice;
- ❖ Sphericity;
- ❖ Compactness 1;
- ❖ Elongation;
- ❖ Surface volume ratio;
- ❖ Volume;
- ❖ Flatness;
- ❖ Spherical disproportion;
- ❖ Roundness;
- ❖ Surface area;
- ❖ Maximum 2D diameter column;
- ❖ Maximum 2D diameter row.

S4 Preprocessing methods for images and data

S4.1 Resampling

In our study, CT images were collected using three different scanners with variable acquisition techniques and parameters; thus, the radiomics were difficult to reproduce. To diminish the influence of the variable CT parameters, we used the resampling method to process the images before extracting features and resample the voxel resolution of all images to 1 mm × 1 mm × 1 mm.

S4.2 Min–Max normalization

Different radiomics features have different value ranges, which makes it difficult to compare two features with variable orders of magnitude. Prior to further analysis, Min–Max normalization was used to normalize the maximum and minimum values in the data column, and the standardized value was between 0 and 1. The following formula was used:

$$X_{norm} = \frac{X - X_{min}}{X_{max} - X_{min}} \quad [1]$$

where X_{norm} , X , X_{min} , and X_{max} represent the normalized data, raw data, minimum value, and maximum value, respectively.

S5. Information of eleven selected radiomics features after dimensionality reduction and feature selection

S5.1 Radiomics features of arterial phase

- ❖ Squareroot_glszm_Size Zone Variability;
- ❖ Original_first order_Median;
- ❖ Original_first order_Total Energy;
- ❖ wavelet-LHH_glcM_Idmn;
- ❖ wavelet-HLL_first order_Median.

S5.2 Radiomics features of venous phase

- ❖ log-sigma-3-0-mm-3D_glcM_Imc2;
- ❖ log-sigma-3-0-mm-3D_glrIm_Run Length Non-Uniformity;
- ❖ Squareroot_first order_Skewness;
- ❖ Squareroot_glcM_Imc2;
- ❖ wavelet-LHL_glszm_Zone Entropy;
- ❖ wavelet-HLL_glszm_Zone Entropy.

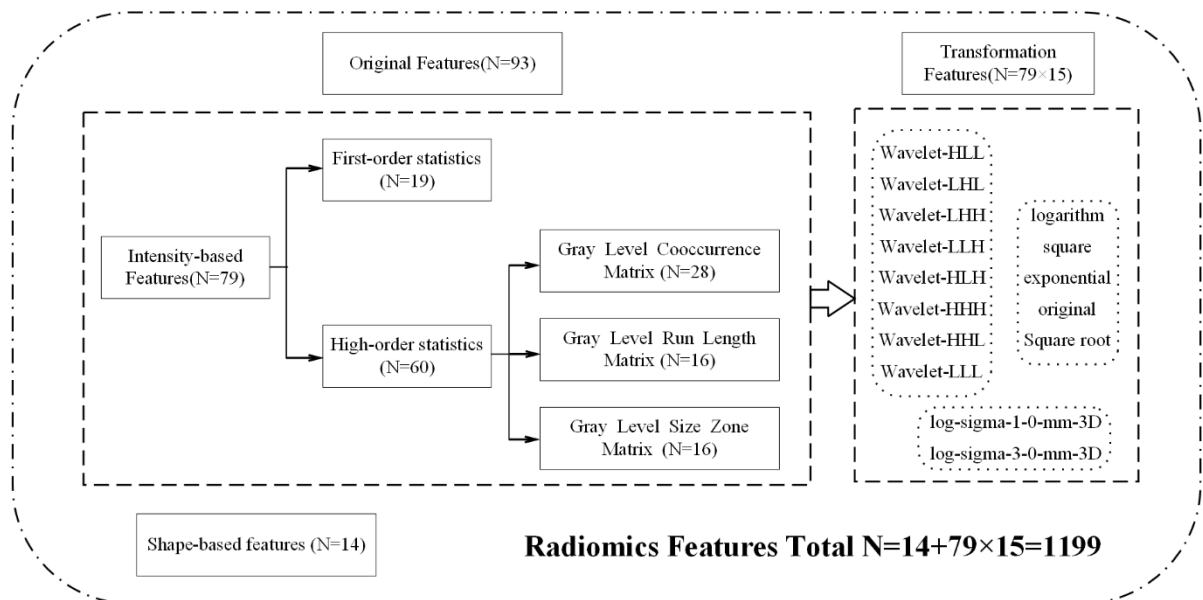


Figure S1 The details of extracted radiomics features.

Table S1 The performance of the radiomics model, clinical model, and combined model in the training cohort, the internal validation cohort and the external validation cohort with the AdaBoost method

Model	Training cohort			Internal validation cohort			External validation cohort		
	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined
AUC (95% CI)	0.709 (0.622, 0.723)	0.902 (0.854, 0.923)	1.000 (0.914, 0.970)	0.603 (0.743, 0.91)	0.913 (0.844, 0.969)	0.869 (0.803, 0.956)	0.595 (0.585, 0.759)	0.837 (0.870, 0.985)	0.788 (0.852, 0.981)
Accuracy	0.741	0.911	1.000	0.679	0.839	0.696	0.670	0.766	0.702
Sensitivity	0.570 (49/86)	0.872 (75/86)	1.000 (86/86)	0.727 (16/22)	0.727 (16/22)	0.955 (21/22)	0.519 (14/27)	0.704 (19/27)	0.667 (18/27)
Specificity	0.848 (117/138)	0.935 (129/138)	1.000 (138/138)	0.647 (22/34)	0.912 (31/34)	0.529 (18/34)	0.731 (49/67)	0.791 (53/67)	0.716 (48/67)
PPV	0.700 (49/70)	0.893 (75/84)	1.000 (86/86)	0.571 (16/28)	0.842 (16/19)	0.568 (21/37)	0.438 (14/32)	0.576 (19/33)	0.486 (18/37)
NPV	0.760 (117/154)	0.921 (129/140)	1.000 (138/138)	0.786 (22/28)	0.838 (31/37)	0.947 (18/19)	0.790 (49/62)	0.869 (53/61)	0.842 (48/57)
F1-score	0.628	0.880	1.000	0.640	0.800	0.712	0.475	0.633	0.562

The brackets in the table show the numerator and denominator of the sensitivity, specificity, PPV and NPV. AUC, area under the receiver operating characteristic curve; CI, confidence interval; PPV, positive predictive value; NPV, negative predictive value.

Table S2 The performance of the radiomics model, clinical model, and combined model in the training cohort, the internal validation cohort and the external validation cohort with the Decision Tree method

Model	Training cohort			Internal validation cohort			External validation cohort		
	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined	Clinical	Radiomics	Combined
AUC (95% CI)	0.5 (0.5, 0.696)	0.821 (0.757, 0.821)	0.813 (0.916, 0.970)	0.5 (0.5, 0.791)	0.766 (0.740, 0.896)	0.735 (0.803, 0.956)	0.5 (0.5, 0.629)	0.695 (0.695, 0.896)	0.688 (0.844, 0.981)
Accuracy	0.616	0.839	0.848	0.607	0.786	0.679	0.713	0.723	0.713
Sensitivity	0.000 (0/86)	0.756 (65/86)	0.663 (57/86)	0.000 (0/22)	0.636 (14/22)	1.000 (22/22)	0.000 (0/27)	0.630 (17/27)	0.630 (17/27)
Specificity	1.000 (138/138)	0.891 (123/138)	0.964 (133/138)	1.000 (34/34)	0.882 (30/34)	0.471 (16/34)	1.000 (67/67)	0.761 (51/67)	0.746 (50/67)
PPV	NaN (0/0)	0.813 (65/80)	0.919 (57/62)	NaN (0/0)	0.778 (14/18)	0.550 (22/40)	NaN (0/0)	0.515 (17/33)	0.500 (17/34)
NPV	0.616 (138/224)	0.854 (123/144)	0.821 (133/162)	0.607 (34/56)	0.789 (30/38)	1.000 (16/16)	0.713 (67/94)	0.836 (51/61)	0.833 (50/60)
F1-score	0.000	0.778	0.770	0.000	0.714	0.710	0.000	0.567	0.557

The brackets in the table show the numerator and denominator of the sensitivity, specificity, PPV and NPV. AUC, area under the receiver operating characteristic curve; CI, confidence interval; PPV, positive predictive value; NPV, negative predictive value; NaN, not a number, which represents an undefined value.