

# Antibiotic exposure windows and the efficacy of immune checkpoint blockers in patients with cancer: a meta-analysis

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**Background:** Immune checkpoint blockers (ICBs) improve the survival of patients with cancer, but primary or acquired drug resistance is inevitable. Intestinal microorganisms play an important role in immunotherapy and antitumor response, and antibiotic use can cause changes in intestinal microbial abundance and diversity. At present, the effects of antibiotic exposure on the anticancer activity of immunotherapy remain controversial.

**Methods:** We performed a meta-analysis of relevant studies retrieved from electronic databases to assess the effects of the time window of antibiotic exposure on the efficacy of immune checkpoint inhibitors (ICIs). In accordance with the definition of antibiotic use in different articles, the time window of antibiotic exposure was divided into three groups, namely, Groups 1 (antibiotic use within 2 months before or after ICI), 2 (antibiotic use before ICI), and 3 (antibiotic use anytime during ICI).

**Results:** After retrieval from the PubMed and the Embase databases, 39 cohorts were included. In group 1, progression-free survival [PFS; hazard ratio (HR) =1.81, 95% confidence interval (CI): 1.40–2.34] and overall survival (OS; HR =1.81, 95% CI: 1.43–2.28) were prolonged in patients without antibiotic use. In group 2, the subgroup analysis showed that antibiotic use had no effect on PFS (HR =0.90, 95% CI: 0.65–1.26) and OS (HR =1.53, 95% CI: 0.89–2.62) when the exposure window defined as 0–3 months. In Group 3, pooled results indicated that PFS (HR =0.78, 95% CI: 0.65–0.93) was prolonged in patients with antibiotic during immunotherapy, and no difference was observed in the OS data (HR =0.98, 95% CI: 0.78–1.24) between the patients with antibiotic and without antibiotic.

**Conclusions:** Antibiotic use in shortly time (within before or after 2 months) around the initiation of immunotherapy was remarkably related to the efficacy of ICIs. A different scenario could be observed that during the long-term treatment of ICIs, the effect of antibiotic exposure seems to be eliminated.

Keywords: Antibiotics; gut microbiome; immune checkpoint inhibitors (ICIs); survival; meta-analysis

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#### Introduction

Immune checkpoint inhibitors (ICIs), including anticytotoxic T-lymphocyte-associated protein 4 (anti-CTLA-4) and anti-programmed cell death protein-(L)-1 [anti-PD-(L)1] monoclonal antibodies (mAbs), reactivate the antitumor activity of CD8<sup>+</sup> T cells by blocking T cell signals and are extensively approved in multiple cancers (1). In recent years, ICIs have dramatically revolutionized the management of multiple types of cancer. Patients with cancer have achieved overall response rates of 13.3–87%, 18–23%, and 11.9–19% by anti-PD-1, anti-PD-L1, and anti-CTLA-4 mAbs, respectively (2). However, some patients with advanced cancer have poor response to ICIs. In this regard, we seek to find the factors that influence the efficacy of ICIs for improved clinical drug use.

The gut microbiome has been demonstrated to affect cancer therapy especially the efficacy of checkpoint inhibitors in patients with advanced cancer (3). Routy et al. have found that the composition of the gut microbiome is different between the responders and the non-responders to ICIs and that fecal microbiota transplantation from responders can improve the efficacy of cancer immunotherapy in non-responders (4). Several retrospective studies have found that the poor efficacy of immunotherapy is associated with antibiotic (ATB) exposure, whereas Hogue et al. (5) have observed the opposite outcome. Also, some studies deny the association. Notably, these studies have not reached a consistent definition on the ATB use especially the time window of ATBs. Thus, we performed a meta-analysis to determine whether the use of ATBs before, during, or after immunotherapy affect the efficacy of ICIs in patients with cancer. This study aimed to explore many predictors for patient with ICIs. We present the following article in accordance with the PRISMA reporting checklist (available at http://dx.doi.org/10.21037/apm-20-2076).

#### Methods

#### Literature search

We conducted a systematic review in the PubMed and the Embase databases by using the terms "(immune checkpoint inhibitor [Title/Abstract]) OR immune checkpoint inhibitors [Title/Abstract]) OR immune checkpoint blockade [Title/Abstract]) OR ICI [Title/Abstract]) OR ICIs [Title/Abstract]) OR ICB [Title/Abstract]) OR immunotherapy [Title/Abstract]) OR immunotherapies [Title/Abstract]" and references from relevant articles in the latest 5 years up to Nov. 7, 2020. The included articles were subjected to a dual review, and the references of the included studies were manually reviewed for any additional publication. We searched the PROSPERO database without restricted and no articles were found. Our registration number was CRD42020155823. As we performed a meta-analysis about researches of published studies, no need application for ethics approval.

#### Quality assessment and data extraction

The data from each study that met the inclusion criteria were independently extracted by two authors (Litang Huang and Xi Chen). Any problem with data extraction was resolved by discussion. The retrieved and the extracted data included the author's name, year of publication, country, study design, cancer types, number of samples (number of patients exposed to ATBs), type of ICIs, ATB window, and outcomes [progression-free survival (PFS)/overall survival (OS), associated hazard ratio (HR), and 95% confidence interval (CI)]. If data were available in both sources, the source with more complete data were prioritized.

#### Grouping

Here, we divided the included studies into three groups in accordance with the time windows of exposure. Group 1 was administered with ATBs within 2 months before or after immunotherapy. Group 2 was injected with ATBs before immunotherapy. Group 3 was exposed to ATBs at any time during the immunotherapy (*Figure 1*).

### Statistical analysis

The survival outcomes, including OS and PFS, were obtained. The effect of the time window of ATB exposure on the survival of patients with immunotherapy was determined using HRs and 95% CIs. Furthermore, the association between ATB exposure window and ICI efficacy was included. A meta-analysis was performed to compute the weighted average of PFS or OS reported for patients with and without exposure to ATB. The  $I^2$  statistic and the P value were used to examine heterogeneity across articles for each outcome. A P value  $\leq 0.05$  was defined as significant heterogeneity. We conducted the subgroup analysis to examine studies in accordance with the type of group (ATB exposure window). The publication bias was assessed using the Begg's test and funnel plots, and significant publication





Figure 2 Literature search and study selection.

bias was defined as P<0.05. All statistical analyses were conducted using the STATA version 15.

#### **Results**

A total of 1,061 relevant reports from the PubMed and the Embase databases were retrieved, and three more studies were identified. A total of 239 studies were removed after duplicate checking, and 723 studies were removed after reviewing the title or the abstract. After screening and eligibility assessment, 99 studies remained for fulltext screening. Sixty-four reports, including 5 reviews, 6 commentaries, 2 meta-analysis, 33 incomplete studies (lacking HR for PFS or OS), and 18 duplications, were subsequently excluded. Three records were identified through meta-analysis. Finally, 39 studies were included in our quantitative analysis (*Figure 2*). Twenty-eight studies were complete cohort studies, whereas the rest was shown only as abstract.

#### **Characteristics**

*Table 1* shows the population distribution and the characteristics of the included studies. A total of 7,853 patients from 39 studies met our inclusion criteria. A total of 2,400 (30.6%) patients were exposed to ATBs. The included studies were published between 2017 and 2020, and most studies were conducted in 2019 (48%). Almost two-fifth (37%) of the studies were from the United States. Of the 39 included studies, 2 and 35 were prospective and

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# Sample Publication Country Study design ATB window Cancer type Immunotherapy (ATB+/ATB-) 74 [15/59] Anti-PD-1 Kaderbhai France Retrospective NSCLC 3 months before or

Table 1 Baseline characteristics

(6) 2017						during ICI	
Thompson (7) 2017	US	Retrospective	NSCLC	74 [18/56]	Anti-PD-1	Within 6 weeks of ICI	$PFS \downarrow / OS \downarrow$
Ahmed (8) 2018	US	Retrospective	Advanced cancer	60 [17/43]	Anti-PD-(L)1	Within 2 weeks before or after ICI	$PFS \downarrow / OS \downarrow$
Chalabi (9) 2020	Switzerland	Retrospective	NSCLC	757 [169/588]	Atezolizumab	Within 1 month before and after ICI	OS↓
Derosa (10)	US	Retrospective	RCC	121 [16/105]	Anti-PD-(L)1 or	1 or 2 months	$PFS{\downarrow}/OS\downarrow$
2018			NSCLC	239 [48/191]	anti-CTLA-4	before ICI	$PFS{\downarrow}/OS\downarrow$
Huemer (11) 2018	Austria	Retrospective	NSCLC	142 [80/62]	Anti-PD-(L)1	1 month before or after ICI	PFS(-)/OS(-)
Kulkarni (12) 2018	US	Retrospective	NSCLC	111 [44/67]	Nivolumab or Pembrolizumab	3 months before ICI	PFS(-)/OS(-)
Kulkarni (13)	US	NR	NSCLC	148 [87/61]	ICI	<1 month or	PFS $\uparrow$ /OS $\uparrow$
2019			RCC	55 [40/15]		during ICI	PFS ↓/OS(–)
Sen (14)	US	Retrospective	RCC/NSCLC ex.	177 [57/120]	Anti-PD-1 or	During ICI	PFS(-)/OS(-)
2018					anti-CTLA-4	Within 1 month of ICI	PFS(–)/OS ↓
						1–2 months before ICI	PFS(-)/OS(-)
Bagley (15) 2019	US	Retrospective	NSCLC/melanoma	NR	ICI	6 weeks before and 4 weeks after ICI	OS(-)
Barrón (16) 2019	US	NR	NSCLC	140 [18/122]	Anti-PD-L1	Within 1 month before ICI	PFS(–)/OS↓
Elkrief (17) 2019	Canada	Retrospective	Melanoma	74 [10/64]	Anti-PD-1 or anti- CTLA-4	1 month before ICI	PFS ↓/OS(–)
Galli (18) 2019	Italy	Retrospective	NSCLC	157 [46/111]	Anti-PD-(L)1 or anti-CTLA-4	1 month before and 3 months after ICI	PFS(-)/OS(-)
Greally (19) 2019	US	Retrospective	Esophagogastric	162 [62/100]	Anti-PD-(L)1 or anti-CTLA-4	During or within 2 months before ICI	PFS(-)/OS(-)
						1 month before ICI	PFS(–)/OS ↓
Hakozaki (20) 2019	Japan	Retrospective	NSCLC	90 [13/77]	Nivolumab	≥3 days within 1m before ICI	PFS ↓/OS(–)
Hogue (5)	US	Retrospective	NSCLC	166 [NR]	Nivolumab or	3 months before ICI	PFS(-)/OS(-)
2019					pembrolizumab	During ICI	PFS↑
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Table 1 (continued)

Outcome

PFS(-)

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Publication	Country	Study design	Cancer type	Sample (ATB+/ATB-)	Immunotherapy	ATB window	Outcome
Kim (21)	Korea	Retrospective	NSCLC/others	234 [108/126]	Anti-PD-(L)1 or	During ICI	PFS(-)
2018					anti-CTLA-4	0–2 months before ICI	$PFS \downarrow / OS \downarrow$
						0–1 month before ICI	$PFS \downarrow / OS \downarrow$
						1–2 months before ICI	PFS ↓/OS(–)
Lalani (22) 2020	US	Retrospective	mRCC	146 [31/115]	Anti-PD-(L)1	8 weeks before and 4 weeks after ICI	PFS ↓/OS(–)
Ouaknine Krief (23) 2019	France	Retrospective	NSCLC	72 [30/42]	Anti-PD-1	2 months before to 1 month after ICI	PFS(–)/OS ↓
Pinato (24)	UK	Prospective	NSCLC/Melanoma/	196 [29/167]	Anti-PD-(L)1	1 month before ICI	OS↓
2019			other			Concurrently to ICI therapy	OS(-)
Schett (25)	Switzerland	Retrospective	NSCLC	218 [33/185]	Anti-PD-(L)1	2 months before ICI	$PFS \downarrow / OS \downarrow$
2019						During ICI	PFS(-)
Tinsley (26) 2020	UK	Retrospective	NSCLC/Melanoma/ mRCC	291 [92/199]	ICI	2 weeks before and 6 weeks after ICI	PFS ↓/OS ↓
Ueda (27) 2019	Japan	Retrospective	RCC	31 [5/26]	Anti-PD-(L)1 or anti-CTLA-4	1 month before ICI	$PFS\downarrow$
Zhao (28) 2019	China	Retrospective	NSCLC	109 [20/89]	Anti-PD-1	1 month before or after ICI	$PFS \downarrow / OS \downarrow$
Routy (4) 2018	France	Retrospective	NSCLC/UC	249 [69/180]	Nivolumab or durvalumab	2 months before or 1 month after ICI	$PFS \downarrow / OS \downarrow$
Agarwal (29) 2019	US	Retrospective	UC	101 [26/75]	Anti-PD-(L)1	1 month before to during ICI	OS↓
Do (30) 2018	US	Retrospective	Lung cancer	109 [87/22]	Anti-PD-1	1 month before to 1 month after ICI	OS↓
Masini (31) 2019	Italy	Retrospective	Advanced cancer	169 [NR]	Nivolumab Pembrolizumab Ipilimumab	During ICI	OS(-)
Rounis (32) 2019	Greece	Prospective	NSCLC	44 [NR]	ICI	1 month before or during ICI	$PFS \downarrow / OS \downarrow$
Guo (33) 2019	China	Retrospective	Esophagogastric	49 [21/18]	ICI	2 months before and 1 month after	$PFS \downarrow / OS \downarrow$
Forde (34) 2020	Ireland	Retrospective	NSCLC	86 [34/52]	ICI	NR	PFS(-)/OS(-)
Hopkins (35) 2020	Australia	Retrospective	UC	896 [235/661]	Atezolizumab	1 month before or after ICI	PFS↓/OS ↓

Table 1 (continued)

Table 1 (continued)

Publication	Country	Study design	Cancer type	Sample (ATB+/ATB-)	Immunotherapy	ATB window	Outcome
Iglesias- Santamaría (36) 2020	Spain	Retrospective	Advanced cancer	102 [60/42]	ICI	4 weeks before or after ICI	PFS(-)/OS(-)
				[33/42]		During ICI	PFS(-)/OS(-)
Kapoor (37) 2020	India	Retrospective	Advanced cancer	155 [70/85]	ICI	2 weeks before ICI	PFS(-)/OS(-)
Khan (38) 2020	NR	Retrospective	UC	146 [31/115]	ICI	2 months before ICI	PFS(-)/OS(-)
				[44/46]		2 months after ICI	PFS(-)/OS(-)
Mohiuddin (39) 2020	NR	Retrospective	Melanoma	568 [114/454]	ICI	3 months before ICI	$OS\downarrow$
Swami (40) 2020	US	Retrospective	Melanoma	166 [30/136]	ICI	Within previous 2 months	PFS↓/OS(–)
Abu-Sbeih	US	Retrospective	Advanced cancer	826 [569/257]	ICI	After ICI	OS↓
(41) 2019						Before ICI	
						Before and after ICI	
Facchinetti (42) 2020	Italy	Retrospective	SCLC	143 [36/107]	Pembrolizumab	2 months before ICI	PFS(-)/OS(-)

Table 1 (continued)

ATB, antibiotic; NSCLC, non-small cell lung cancer; anti-PD-(L)1, anti-programmed cell death protein-(L)-1; PFS, progression-free survival; OS, overall survival; RCC, renal cell carcinoma; ICI, immune checkpoint inhibitor; anti-CTLA-4, anti-cytotoxic T-lymphocyte-associated protein 4; NR, not reported; US, United States; UC, urothelium carcinoma; UK, United Kingdom.

retrospective studies, and two studies did not mention the type of study. The patients were diagnosed with lung cancer (49%), renal cell carcinoma (about 6%), melanoma (about 13%) and other advanced cancers, including esophageal cancer and urothelium carcinoma. The ICIs included anti-PD-(L)1 and anti-CTLA-4. The ATB window had different definitions in the studies (*Table 1*).

#### Outcome data

#### Survival of group 1

Group 1 included 3,237 patients from 14 studies. These patients mostly had non-small cell lung and urethral cancers. Pooled results showed that the ATB exposure were negatively associated with the PFS (HR =1.81, 95% CI: 1.40–2.34,  $I^2$ =55.0%) and the OS (HR =1.81, 95% CI: 1.43–2.28,  $I^2$ =61.5%) of patients who underwent immunotherapy (*Figure 3*). The PFS and the OS were analyzed using the random-effects models due to significant heterogeneity.

#### Survival of group 2

Group 2 was divided into three subgroups on the basis of the duration of the ATB exposure before immunotherapy. Subgroups 1, 2, and 3 were exposed to ATB before immunotherapy within 1, 2, and 3 months, respectively. The pooled results of subgroups 1 and 2 showed that ATB was a risk factor of poor OS (subgroup 1: HR =2.25, 95% CI: 1.42–3.55; subgroup 2: HR =1.57, 95% CI: 1.16–2.11) and PFS (HR =1.70, 95% CI: 1.35–2.14; subgroup 2: HR =1.45, 95% CI: 1.04–2.02). However, the results of subgroup 3 showed that the ATB use was not related to the OS (HR =1.53, 95% CI: 0.89–2.62) and the PFS (HR =0.90, 95% CI: 0.65–1.26) of patients with cancer who received immunotherapy. In subgroup 3, two cohorts for OS data and two cohorts for PFS data were available (*Figure 4*).

#### Survival of group 3

Four cohorts were included for analysis. Pooled results showed that ATB use could prolong the PFS (HR =0.78,



Figure 3 The associations between antibiotic exposure and PFS (A) and OS (B) in group 1. ES, effect size; CI, confidence interval; PFS, progression-free survival; OS, overall survival.

В

Study	%
ID 0~1 month	ES (95% CI) Weigh
Hakozaki.2018	2.69 (1.26, 5.77) 7.58 2.20 (1.30, 3.30) 15.52
Derosa.2018	1.30 (0.90, 1.80) 21.57
Elkrief.2019	3.13 (1.20, 7.69) 5.41
Sen.2018	1.20 (0.80, 2.10) 14.84
Ueda.2019	
Subtotal (I-squared = 31.8%, p = 0.174)	1.70 (1.35, 2.14) 100.0
0~2 month Derosa.2018	3.20 (1.60, 5.90) 11.92
Derosa.2018	1.20 (0.90, 1.60) 18.98
Kim.2018	1.72 (1.26, 2.33) 18.65 3 45 (1 44, 8 29) 8 70
Khan.2020	0.91 (0.40, 1.79) 10.42
Swami.2020	1.28 (0.80, 2.04) 15.36
Subtotal (I-squared = 71.0%, p = 0.002)	1.45 (1.04, 2.02) 100.0
0~3 month	1 02 /0 64 1 62 60 00
Kulkarni.2018	0.80 (0.50, 1.29) 49.31
Subtotal (I-squared = 0.0%, p = 0.474)	0.90 (0.65, 1.26) 100.0
NOTE: Weights are from random effects analysis	
.0786 1	12.7
Study ID	% ES (95% CI) Weigh
0~1 month	2 02 (0 70 5 82) 8 20
Hakuzaki.zu to	2.02 (0.70, 5.03) 0.20
Derosa.2018	2.10 (0.90, 5.00) 9.68
Derosa.2018	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.00
Derosa.2018   Derosa.2018   Kim.2018   Ekrief.2019	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84
Derosa.2018 Derosa.2018 Elixief.2019 Pinato.2019 Barron.2019 Sen.2018	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000)	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Subtotal (I-squared = 89.0%, p = 0.000) 0~2 month Derosa.2018	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) O~2 month Derosa.2018 Derosa.2018	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) 0~2 month Derosa.2018 Kim.2018 C-abert 2010	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 2.72 (4.24, 4.20) 20.33 2.78 (4.24, 4.20) 20.33 3.78 (4.24, 4.20) 20.
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) - 0~2 month Derosa.2018 Kim.2018 Center State Sta	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.8 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 3.73 (1.34, 10.40) 6.60 1.46 (0.65, 3.28) 9.31
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) 0~2 month Derosa.2018 Kim.2018 Schett.2019 Khan.2020 Swami.2020	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.8 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 3.73 (1.34, 10.40) 6.60 1.46 (0.65, 3.28) 9.31 1.73 (1.00, 2.99) 14.80
Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) O~2 month Derosa.2018 Kim.2018 Schett.2019 Khan.2020 Swami.2020 Swami.2020 Subtotal (I-squared = 53.2%, p = 0.046)	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 3.73 (1.34, 10.40) 6.60 1.46 (0.65, 3.28) 9.31 1.73 (1.00, 2.99) 14.80 0.80 (0.50, 1.20) 18.04 1.57 (1.16, 2.11) 100.0
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Derosa.2018 Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) 0~2 month Derosa.2018 Kim.2018 Schett.2019 Khan.2020 Swami.2020 Swami.2020 Swami.2020 Subtotal (I-squared = 53.2%, p = 0.046) 0~3 month Hogue.2019 Mohiuddin.2020 Subtotal (I-squared = 71.9%, p = 0.059)	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 3.73 (1.34, 10.40) 6.60 1.46 (0.65, 3.28) 9.31 1.73 (1.00, 2.99) 14.80 0.80 (0.50, 1.20) 18.04 1.57 (1.16, 2.11) 100.0 1.12 (0.69, 1.82) 44.12 1.95 (1.43, 2.66) 55.88 1.53 (0.89, 2.62) 100.0
Derosa.2018 Kim.2018 Elkrief.2019 Pinato.2019 Barron.2019 Sen.2018 Kapoor.2020 Subtotal (I-squared = 89.0%, p = 0.000) 0~2 month Derosa.2018 Schett.2019 Khan.2020 Swami.2020 Subtotal (I-squared = 53.2%, p = 0.046) 0~3 month Hogue.2019 Mohiuddin.2020 Subtotal (I-squared = 71.9%, p = 0.059) NOTE: Weights are from random effects analysis	2.10 (0.90, 5.00) 9.68 2.50 (1.60, 3.70) 13.06 2.27 (1.50, 3.43) 13.10 2.00 (0.83, 4.76) 9.55 3.40 (1.90, 6.10) 11.84 → 7.80 (2.40, 25.00) 7.47 2.00 (1.20, 3.30) 12.43 1.02 (1.00, 1.04) 14.66 2.25 (1.42, 3.55) 100.0 2.00 (0.90, 4.30) 9.76 1.60 (1.10, 2.30) 20.33 1.78 (1.26, 2.52) 21.17 3.73 (1.34, 10.40) 6.60 1.46 (0.65, 3.28) 9.31 1.73 (1.00, 2.99) 14.80 0.80 (0.50, 1.20) 18.04 1.57 (1.16, 2.11) 100.0 1.12 (0.69, 1.82) 44.12 1.95 (1.43, 2.66) 55.88 1.53 (0.89, 2.62) 100.0

Figure 4 The associations between antibiotic exposure and PFS (A) and OS (B) in group 2. ES, effect size; CI, confidence interval; PFS, progression-free survival; OS, overall survival.

95% CI: 0.65–0.93) during immunotherapy. By contrast, the ATB use during immunotherapy was not related to the OS (HR =0.98, 95% CI: 0.78–1.24) of patients with cancer (*Figure 5*).

#### Publication bias analysis

The Begger's funnel plot was used to assess the publication bias in this meta-analysis. Results indicated no publication bias in any study, as evidenced by the symmetrical funnel plots (Figures S1–S3).

#### Discussion

Our meta-analysis has reported the relationship between the ATB exposure window and the efficacy of ICIs in patients with cancer. However, in the published metaanalysis, different results on the effect of ATBs on ICIs are observed. Huang et al. believe that ATB use was associated with poor survival in patients with immunotherapy (43). However, Wilson et al. have found that when a very broad definition of antibiotic exposure is adopted (antibiotic exposure anytime within the window 60 days before anytime after initiation of immunotherapy), the negative effect of antibiotic to PFS and OS was eliminated (44). Based on the work of Wilson et al., we have re-divided the included cohorts into three groups in accordance with the different definitions of the ATB exposure window to avoid the overlapping definitions of ATB time in different studies as much as possible. We have investigated the effects of ATB exposure on the antitumor efficacy and the survival of ICIs during immunotherapy. Group 1 (ATB use within 2 months before or after ICI) indicates that ATB use is a prognostic factor in immunotherapy. In group 2 (ATB use before ICI), the subgroup analysis shows that ATB use has no effect on immunotherapy when the exposure window is defined as 0-3 months. Although the ATB exposure window of the patients included cannot be completely distinguished, the cohorts have no detailed data about the patients exposed. The prolonged time between the exposure of ATB and the start of ICI may lead to the disappearance of adverse prognosis caused by ATB. Many studies suggest that ATBs may cause the poor efficacy of immunotherapy by affecting the abundance or imbalance of intestinal flora, and the gut flora may return to baseline after 42 days (44). In group 3, pooled results show that ATB exposure is positively correlated with the PFS but not with the OS, Tinsley et al. noted that retrospective studies which failed to show any

association between antibiotic therapy and ICI efficacy (26). Facchinetti *et al.* found that Eastern Cooperative Oncology Group performance status (ECOG PS) 2 was the only factor independently impacting on both PFS and OS (42), even though Hopkins *et al.* found the negative impact of antibiotic exposure, but the authors themselves are cautious in their interpretation of results, with a special situation was detected that ECOG PS was generally low in the cohort (35). Our results suggested that during the treatment of ICI, if ATB are required, perhaps, it may not cause the negative impact of efficacy of ICI. As we all known, patients with infection may cause bad PS, the findings about negative impact of antibiotic use which may be confounded by overall health status of patients that necessitates antibiotic use.

The limitations of our study are the same as those of several other published meta-analyses. The included studies are retrospective studies. Although we classify ATB exposure windows as best as we can, an overlap remains. In addition, the lack of baseline characteristics of the included patients, such as the type of ATB, specific infection site, duration of ATB use, and PS of patients, has prevented further subgroup analysis.

Several studies classify the patients who received ICIs into responders and non-responders in accordance with the best clinical response as assessed by the RECIST1.1 (4,45). The baseline gut microbiome diversity and the relative abundance of the two groups are different, as shown by the higher relative abundance of the Akkermansia of the responder. The fecal microbiota of the two groups of patients are transplanted to specific pathogen-free mice. The mice transplanted with the microbiota of nonresponders had inferior response to ICI. Patients with high gut microbiome diversity and high relative abundance of some symbiotic bacteria are likely to benefit from ICIs. Studies have shown that the use of ATBs can affect the intestinal microbial diversity, thereby affecting the efficacy of ICIs. Different types of ATBs have different effects on the gut microbiome function. Mohiuddin et al. have found that the response of patients to ICIs is affected by the type of ATBs they use. Penicillin has the most serious adverse effects followed by cephalosporins and quinolones. However, vancomycin has no effect on the survival of patients (39). This article includes retrospective studies and cannot obtain the specific baseline characteristics of the included patients. Among all patients receiving immunotherapy, most patients using ATBs have respiratory or urinary tract infections. The immune characteristics,



Figure 5 The associations between antibiotic exposure and PFS (A) and OS (B) in group 3. ES, effect size; CI, confidence interval; PFS, progression-free survival; OS, overall survival.

baseline intestinal microbial, and ECOG status of patients with ATB exposure are different from those without ATB exposure. In some cases, such as patients with bacteremia, the use of ATBs is inevitable. ATBs improve the response of such patients to immunosuppressants by inhibiting pathogenic bacteria, and this finding may partly explain the high PFS of patients taking ATBs in group 3 (ATB use at any time during the ICIs). Therefore, we need to understand the baseline characteristics of patients using ATBs and the dynamic changes in their intestinal microbes after using different ATBs. Summarizing from the current research data, high-dose broad-spectrum ATBs (such as cephalosporins,  $\beta$ -lactams, and quinolones) may affect the intestinal flora, impair the efficacy of immunotherapy, and shorten the survival time of patients (reviewed and nonprospective data). The timing of ATBs is important. Before immunotherapy, if infections are present, the corresponding anti-infective treatment based on bacteriological evidence is recommended to be provided to avoid the prophylactic and the long-term use of ATBs.

#### Conclusions

This meta-analysis included 30 cohorts. Results showed that the survival of patients with cancer who underwent immunotherapy was associated with ATB exposure and that the timing of ATB use was an important factor. Different ATB exposure windows had different effects on the survival of patients with cancer. In the future, advanced prospective studies are needed to guide immunotherapy accurately and improve the patients' survival.

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#### Footnote

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*Conflicts of Interest*: All authors have completed the ICMJE uniform disclosure form (available at http://dx.doi. org/10.21037/apm-20-2076). The authors have no conflicts of interest to declare.

*Ethical Statement*: The authors are accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

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Begg's funnel plot with pseudo 95% confidence limits



Figure S1 Funnel plot of overall survival in group 1.



Figure S2 Funnel plot of progression-free survival in group 2.

Begg's funnel plot with pseudo 95% confidence limits



Figure S3 Funnel plot of overall survival in group 2.