

# Preoperative nomogram to predict survival following colorectal cancer liver metastasis simultaneous resection

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**Background:** Simultaneous resection for patients with synchronous colorectal cancer liver metastases (CRLM) remains an optimal option for the sake of curability. However, few studies so far focus on outcome of this subgroup of patients (who receive simultaneous resection for CRLM). Substantial heterogeneity exists among such patients and more precise categorization is needed preoperatively to identify those who may benefit more from surgery. In this study, we formulated this internally validated scoring system as an option. **Methods:** Clinicopathological and follow-up data of 234 eligible CRLM patients undergoing simultaneous resection from January 2010 to March 2019 in our center were included for analysis. Patients were randomized to either a training or validation cohort. We performed multivariable Cox regression analysis to determine preoperative factors with prognostic significance using data in training cohort, and a nomogram scoring system was thus established. Time-dependent receiver operating characteristic (ROC) curve and calibration plot were adopted to evaluate the predictive power of our risk model.

**Results:** In the multivariable Cox regression analysis, five factors including presence of node-positive primary defined by enhanced CT/MR, preoperative CEA level, primary tumor location, tumor grade and number of liver metastases were identified as independent prognostic indicators of overall survival (OS) and adopted to formulate the nomogram. In the training cohort, calibration plot graphically showed good fitness between estimated and actual 1- and 3-year OS. Time-dependent ROC curve by Kaplan-Meier method showed that our nomogram model was superior to widely used *Fong's* score in prediction of 1- and 3-year OS (AUC 0.702 *vs.* 0.591 and 0.848 *vs.* 0.801 for 1- and 3-year prediction in validation cohort, respectively). Kaplan-Meier curves for patients stratified by the assessment of nomogram showed great discriminability (P<0.001).

**Conclusions:** In this retrospective analysis we identified several preoperative factors affecting survival of synchronous CRLM patients undergoing simultaneous resection. We also constructed and validated a risk model which showed high accuracy in predicting 1- and 3-year survival after surgery. Our risk model is expected to serve as a predictive tool for CRLM patients receiving simultaneous resection and assist physicians to make treatment decision.

Keywords: Colorectal cancer (CRC); liver metastasis; nomogram; survival; preoperative prediction

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#### Introduction

Colorectal cancer (CRC) is the fourth most prevalent malignancy and the fifth leading cause of cancer death worldwide (1). Liver is the most common site of metastases from CRC and approximately 15% of colorectal cancer liver metastases (CRLM) present liver lesion at the initial diagnosis (synchronous metastasis) (2,3). Despite advances in systemic therapy, surgery is still an irreplaceable part of modern therapy approach for CRLM. The median survival of CRLM without treatment is 9.6 months, while the 5-year overall survival (OS) of resected CRLM approaches 50 months (4). In recent years, surgical strategies of synchronous CRLM have changed dramatically. Although several studies revealed the surgical outcome or survival advantage among classical colorectum-first, liver-first, and synchronous surgery is similar for CRLM, synchronous resection has gradually been accepted and popularized with advantages of safety, minimal damage, and less cost (5-8). However, few studies so far have focused on outcome of this subgroup of patients who receive simultaneous resection for CRLM. The survival of CRLM patients undergoing simultaneous resection remains highly variable and more precise categorization is needed preoperatively to identify those who may benefit more from surgery. Fong's score (9) is a well-accepted prognostic system initially formulated for metachronous CRLM, and GAME score (10), put forward by Margonis, is a preoperative model based on American cohort requiring genetic status of KRAS, which is not specifically designed for simultaneous CRLM resection. Therefore, we provided this scoring system merely requiring clinicopathological data as an option for CRLM patients undergoing simultaneous resection for preoperative risk stratification.

The present study assessed the preoperative prognostic factors in CRLM patients undergoing synchronous resection. A nomogram was established and validated to quantify the impact of every variable. We present this article in accordance with the TRIPOD reporting checklist (11) (available at http://dx.doi.org/10.21037/jgo-20-329).

#### Methods

#### Patients and data sources

Clinicopathological information of 415 consecutive patients with CRLM who underwent combined liver and colon/ rectum resection with or without chemoradiotherapy from January 2010 to March 2019 in Changhai Hospital was retrospectively collected for analysis. Telephone followup was completed by May 2020. Inclusion criteria were as follows: (I) verified diagnosis of CRC by colonoscope and biopsy; (II) synchronous liver metastasis demonstrated by enhanced computed tomography (CT) scan of abdomen or enhanced liver magnetic resonance (MR); (III) R0 resection of primary lesion and liver metastasis; (IV) definite diagnosis of CRC with liver metastasis by postoperative biopsy; (V) treated with postoperative systemic chemotherapy (5-FUbased) ± targeted therapy. Exclusion criteria were as follows: (I) presence of extrahepatic metastasis; (II) 30-day mortality from operative complications; (III) loss to follow-up or absence of clinicopathological information. A total of 234 cases meeting the eligibility criteria above were included in this study.

Preoperative assessment consisted of gender, age, primary tumor location, preoperative serum CEA and CA-199, number and diameter of liver metastases, presence of nodepositive primary defined by enhanced CT/MR, tumor grade obtained via endoscopic biopsy, and neoadjuvant therapy, all transformed as categorical or hierarchical variables. Postoperative pathological characteristics including primary tumor location, tumor grade, pT stage, pN stage, tumor deposit (TD, focal aggregates of cancer cells in the pericolic and mesenteric adipose tissue around primary tumor). perineural invasion (PNI), lymphovascular invasion (LVI), status of KRAS (exon 2, 3 and 4), NRAS (exon 2, 3 and 4) and BRAF (V600E), number and maximal diameter of liver metastases was obtained. All the auxiliary examination findings were interpreted by experienced specialists unaware of patients' clinical condition. The American Joint Committee on Cancer (AJCC) 7th edition was adopted in most cases and pathologists were consulted to ensure consistency of criteria (12). Simultaneous hepatectomy was performed by experienced hepatologic surgeons who assessed imaging finding to determine surgical procedure. Major liver resection was defined as resection of three or more segments. Survival information was obtained via telephone follow-up survey by May 2020.

The study was conducted in accordance with the Declaration of Helsinki (as revised in 2013). This study was approved by the Institutional Review Board of Changhai Hospital, Secondary Military Medical University, Shanghai, China (No. CHEC2015-146). Because of the retrospective nature of the study, the requirement for informed consent was waived.



Figure 1 Flowchart of case selection and distribution in this study.

#### Statistical analysis

Categorical variables were analyzed with non-parametric test and continuous data with Student *t*-test or Log-rank test. Univariable and multivariable Cox regression analyses were employed to assess prognostic factors. Survival analyses were performed with Kaplan-Meier method, and compared by log-rank test.

The 234 patients were grouped into training cohort and validation cohort by a ratio of 3 to 1 (174:60) with random number (Figure 1). Univariable and multivariable Cox regression analyses to identify predictors of survival were performed using data in training cohort. For the sake of comprehensive inspection of relevant factors, variables with P<0.15 in univariable analysis were entered into multivariable analysis. Hazard ratio (HR) and 95% confidence interval were reported and variables with P less than 0.05 in multivariable analysis were considered statistically significant. Based on the results of multivariable Cox regression analysis, a nomogram was generated to predict 1- and 3-year OS of CRLM patients undergoing synchronous resection. The predictive accuracy and discriminative ability were evaluated by calibration plot and survival curve, respectively. We adopted time-dependent receiver operating characteristic (ROC) curve to compare the accuracy of the nomogram with Fong's clinical risk

score (CRS) system (9).

The demographical comparison was completed with SPSS 22.0 (SPSS, Inc., Chicago, IL, USA), and Cox proportional hazard model, nomogram, time-dependent ROC, calibration plot and Kaplan-Meier curve with R 3.6.2.

#### Results

#### Clinicopathological characteristics of the patients

Among these 234 patients, 174 patients were included in the training cohort while 60 in the validation cohort. The 3-year OS in this study was 30.3% and median followup time was 26 months. Preoperative clinicopathological characteristics comparison in the two groups showed no statistical differences ( $P \ge 0.05$ , *Table 1*). Also, survival analysis comparing OS between patients receiving surgery in different periods showed stable procedure outcome through the follow-up (Figure S1, P=0.47).

#### Preoperative prognostic factors of OS in the training cohort

In training cohort, the hazard ratio (HR) and P value with 95% confidence interval of every candidate predictor generated by Cox regression analysis are shown in *Table 2*. Presence of positive-node primary defined by enhanced

| Variables                        | All, No. (%) | Training, No. (%) | Validation, No. (%) | Р     |
|----------------------------------|--------------|-------------------|---------------------|-------|
| Subjects, n                      | 234          | 174               | 60                  |       |
| Gender                           |              |                   |                     | 0.058 |
| Male                             | 126 (53.85)  | 100 (57.47)       | 26 (43.33)          |       |
| Female                           | 108 (46.15)  | 74 (42.53)        | 34 (56.67)          |       |
| Age                              |              |                   |                     | 0.947 |
| <65 years old                    | 164 (70.09)  | 122 (70.11)       | 42 (70.00)          |       |
| ≥65 years old                    | 70 (29.91)   | 52 (29.89)        | 18 (30.00)          |       |
| Primary tumor location           |              |                   |                     | 0.768 |
| Left-sided                       | 148 (63.25)  | 111 (63.79)       | 37 (61.67)          |       |
| Right-sided                      | 86 (36.75)   | 63 (36.21)        | 23 (38.33)          |       |
| Preoperative CEA, ng/mL          |              |                   |                     | 0.541 |
| ≤100                             | 205 (87.61)  | 150 (86.21)       | 55 (91.67)          |       |
| >100 and ≤200                    | 17 (7.26)    | 14 (8.05)         | 3 (5.00)            |       |
| >200                             | 12 (5.13)    | 10 (5.75)         | 2 (3.33)            |       |
| Preoperative CA-199              |              |                   |                     | 0.617 |
| ≤200                             | 194 (82.91)  | 143 (82.18)       | 51 (85.00)          |       |
| >200                             | 40 (17.09)   | 31 (17.82)        | 9 (15.00)           |       |
| No. of metastases                |              |                   |                     | 0.651 |
| 1                                | 119 (50.85)  | 90 (51.72)        | 29 (48.33)          |       |
| >1                               | 115 (49.15)  | 84 (48.28)        | 31 (51.67)          |       |
| Diameter of metastasis           |              |                   |                     | 1.000 |
| <5                               | 195 (83.33)  | 145 (83.33)       | 50 (83.33)          |       |
| ≥5                               | 39 (16.67)   | 29 (16.67)        | 10 (16.67)          |       |
| Histologic type                  |              |                   |                     | 0.514 |
| Highly/moderately differentiated | 205 (87.61)  | 151 (86.78)       | 54 (90.00)          |       |
| Poorly differentiated/mucinous   | 29 (12.39)   | 23 (13.22)        | 6 (10.00)           |       |
| Nodal positivity by imaging      |              |                   |                     | 0.691 |
| Negative                         | 79 (33.76)   | 60 (34.48)        | 19 (31.67)          |       |
| Positive                         | 155 (66.24)  | 114 (65.52)       | 41 (68.33)          |       |
| Neoadjuvant therapy              |              |                   |                     | 0.240 |
| Negative                         | 166 (70.94)  | 127 (72.99)       | 39 (65.00)          |       |
| Positive                         | 68 (29.06)   | 47 (27.01)        | 21 (35.00)          |       |
| Procedure of hepatic resection   |              |                   |                     | 0.271 |
| Major resection                  | 88           | 69                | 19                  |       |
| Minor resection                  | 146          | 105               | 41                  |       |
| Transfusion                      |              |                   |                     | 0.360 |
| Yes                              | 109          | 78                | 31                  |       |
| No                               | 125          | 96                | 29                  |       |

CA-199, carbohydrate antigen-199; CEA, carcino-embryonic antigen.

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Table 2 Univariable and multivariable analysis in the training cohort

| Verieblee                        | Univariable         | Univariable |                     | Multivariable |  |
|----------------------------------|---------------------|-------------|---------------------|---------------|--|
| variables                        | HR (95% CI)         | Р           | HR (95% CI)         | Р             |  |
| Gender                           |                     | 0.568       |                     |               |  |
| Male                             | 1                   |             |                     |               |  |
| Female                           | 0.869 (0.538–1.406) |             |                     |               |  |
| Age, years old                   |                     | 0.474       |                     |               |  |
| <65                              | 1                   |             |                     |               |  |
| ≥65                              | 0.821 (0.479–1.409) |             |                     |               |  |
| Primary tumor location           |                     | 0.007       |                     | 0.024*        |  |
| Left-sided                       | 1                   |             | 1                   |               |  |
| Right-sided                      | 1.928 (1.201–3.095) |             | 1.769 (1.077–2.904) |               |  |
| Preoperative CEA, ng/mL          |                     |             |                     |               |  |
| <100                             | 1                   |             | 1                   |               |  |
| 100–200                          | 2.103 (1.029–4.300) | 0.042       | 1.896 (0.919–3.912) | 0.084         |  |
| ≥200                             | 2.203 (1.000–4.857) | 0.050       | 2.287 (1.020–5.129) | 0.045*        |  |
| Preoperative CA-199, ng/mL       |                     | 0.655       |                     |               |  |
| <200                             | 1                   |             |                     |               |  |
| ≥200                             | 0.863 (0.452–1.646) |             |                     |               |  |
| No. of metastases                |                     | 0.058       |                     | 0.007*        |  |
| 1                                | 1                   |             | 1                   |               |  |
| >1                               | 1.588 (0.985–2.560) |             | 1.991 (1.207–3.283) |               |  |
| Diameter of Metastasis, cm       |                     | 0.240       |                     |               |  |
| <5                               | 1                   |             |                     |               |  |
| ≥5                               | 1.387 (0.804–2.393) |             |                     |               |  |
| Tumor grade                      |                     | 0.111       |                     | 0.046*        |  |
| Highly/moderately differentiated | 1                   |             | 1                   |               |  |
| Poorly differentiated/mucinous   | 1.698 (0.886–3.254) |             | 2.009 (1.011–3.992) |               |  |
| Positive lymph node by imaging   |                     | 0.002       |                     | 0.001*        |  |
| Negative                         | 1                   |             | 1                   |               |  |
| Positive                         | 2.434 (1.370–4.327) |             | 2.598 (1.442–4.681) |               |  |
| Neoadjuvant therapy              |                     | 0.194       |                     |               |  |
| Negative                         | 1                   |             |                     |               |  |
| Positive                         | 0.660 (0.352–1.236) |             |                     |               |  |

The symbol "\*" denotes P less than 0.05 in multivariable analysis. CA-199, carbohydrate antigen-199; CEA, carcino-embryonic antigen.



Figure 2 Nomogram based on Cox proportional hazard model to quantify risk of death using five indicators, also shown here are relationship between total points and probabilities of 1- and 3-year survival. CEA, carcino-embryonic antigen.

CT/MR, preoperative CEA level and primary tumor site are statistically associated with OS in univariable analysis. These three factors combined with tumor grade (P=0.111) and number of liver metastases (P=0.058) were included in further multivariable analysis and all these five variables were demonstrated to be independent risk factors associated with OS of patients receiving simultaneous resection for CRLM. Patients with higher preoperative CEA level, right-sided primary location, poor tumor grade, positive lymph node and multiple liver metastasis tend to have less opportunity of long-term postoperative survival according to above results.

#### Generation and validation of predictive nomogram for OS

A nomogram was formulated based on five independent risk factors selected by multivariable analysis of Cox proportional hazard model to visualize and quantify the weight of every factor (*Figure 2*). We applied the nomogram to all cases in both training and validation cohorts and calculated the total score and predicted probability of 1- and 3-year survival of every patient. Calibration plot graphically showed good fitness between estimated and actual 1- and 3-year survival in both cohorts (*Figure 3*).

# Evaluation of the preoperative prognostic nomogram model

Time-dependent ROC curve by Kaplan-Meier method was adopted to assess the predictive power. To estimate the predictive validity of our scoring system, we applied modified Fong's scoring system: (I) dissected lymph node biopsy replaced with MR finding, and (II) disease-free interval uniformly deemed as less than 12 months. The result showed that our nomogram scoring system can better predict 1-year survival [area under curve (AUC) 0.788 vs. 0.652 and 0.702 vs. 0.591 in training and validation cohort respectively]. As for 3-year prediction, modified Fong's score retained high accuracy, with AUC 0.712 and 0.801 in training and validation cohorts respectively, while slightly lower than those of our nomogram (0.752 and 0.848, Figure 4). However, in predicting 5-year OS, both Fong's score and our model showed low power of test, partially reflecting the heterogeneity among long-surviving CRLM patients.

For convenience of clinical use, we stratified patients with maxstat, an algorithm in R used to decide on cutpoint to yield the most significant difference. Based on the result, a total mark of 135 was identified as the cutpoint and patients were divided into low- and high-risk groups. Survival analysis was performed by Kaplan-Meier



Figure 3 Calibration curve showing fitness between predicted and actual survival (the left-sided two plots for training cohort and right-sided for validation cohort).

method with log-rank test and graphically displayed in *Figure 5*, showing good discriminability of this stratification solution (P<0.001). Among three groups, there were statistical differences in several postoperative pathological characteristics including primary location, tumor grade, pT stage, pN stage, peripheral nervous invasion, and number of liver metastases (P<0.05, *Table 3*).

#### Discussion

Combined primary tumor and liver resection is the most preferred curative treatment for synchronous CRLM and should be taken into consideration when R0 resection is possible (13). To date, simultaneous surgery has been accepted widely for its perioperative safety, good long-term effects and improved economic efficiency (14,15). However, when CRLM patients decide to receive synchronous surgery for complete R0 resection, the individual oncological benefit is still difficult to predict. Therefore, optimal preoperative prognostic models for patient selection are required.

The present study demonstrated that preoperative factors including primary tumor location, preoperative CEA level, number of liver metastases, tumor grade and positive lymph node by imaging were independent indicators to predict survival benefit from surgery, which was in line with findings of certain reports (16-18). Notably, there



Figure 4 Time-dependent receiver operating curve comparing power of test between Fong's and our scoring system. AUC, area under curve; CRS, clinical risk score; FP, false positivity; KM, Kaplan-Meier; TP, true positivity.

was no significant difference in OS between the patients receiving neoadjuvant chemoradiotherapy (nCRT) or not. Despite that the National Comprehensive Cancer Network (NCCN) recommends 6 months of perioperative chemotherapy, standard treatment strategy is poorly defined (19). Neoadjuvant therapy has been proposed to be an essential part of comprehensive treatment, and individual response to nCRT reflects tumor biologic characteristic and prognosis (20,21), while redundant chemoradiotherapy has been reported to cause liver damage and poor short-term outcome (22). Our data suggested that neoadjuvant therapy was not essential when patients with resectable CRLM treated with postoperative chemotherapy  $\pm$  targeted therapy.

Nomograms have been considered a reliable tool to quantify risk factors of prognosis (23,24). In our study, a nomogram based on five prognostic indicators was generated to predict 1- and 3-year OS of synchronous CRLM patients undergoing combined resection. The predictive accuracy and the discriminative ability of the nomogram were internally and externally validated, and the results showed good fitness between estimated and actual 1- and 3-year OS. At present, Basingstoke score system, Nordlinger scoring system, Iwatsuki scoring system and Fong's CRS scoring system, which requires postoperative



Figure 5 Survival analysis comparing overall survival of patients stratified by nomogram score, using Kaplan-Meier method and Log-rank test.

 Table 3 Clinicopathologic comparison among different risk groups

 defined by nomogram score

| Variables                        | Low-risk,<br>n=108 | High-risk,<br>n=126 | Ρ       |
|----------------------------------|--------------------|---------------------|---------|
| Primary location                 |                    |                     |         |
| Left-sided or rectum             | 84                 | 64                  | <0.001* |
| Right-sided                      | 24                 | 62                  |         |
| Tumor grade                      |                    |                     |         |
| Highly/moderately differentiated | 104                | 101                 | <0.001* |
| Poorly differentiated/ mucinous  | 4                  | 25                  |         |
| pT stage                         |                    |                     |         |
| 1                                | 1                  | 0                   | 0.008*  |
| 2                                | 7                  | 3                   |         |
| 3                                | 92                 | 102                 |         |
| 4                                | 8                  | 21                  |         |
| pN stage                         |                    |                     |         |
| 0                                | 60                 | 18                  | <0.001* |
| 1                                | 31                 | 67                  |         |
| 2                                | 17                 | 41                  |         |

| Table 3 (continued)        |                    |                     |         |
|----------------------------|--------------------|---------------------|---------|
| Variables                  | Low-risk,<br>n=108 | High-risk,<br>n=126 | Ρ       |
| Tumor deposit              |                    |                     |         |
| Negative                   | 79                 | 78                  | 0.068   |
| Positive                   | 29                 | 48                  |         |
| PNI                        |                    |                     |         |
| Negative                   | 88                 | 87                  | 0.029*  |
| Positive                   | 20                 | 39                  |         |
| LVI                        |                    |                     |         |
| Negative                   | 87                 | 96                  | 0.420   |
| Positive                   | 21                 | 30                  |         |
| KRAS                       |                    |                     |         |
| Wild type                  | 53                 | 51                  | 0.094   |
| Mutant type                | 37                 | 39                  |         |
| Information absent         | 18                 | 36                  |         |
| NRAS                       |                    |                     |         |
| Wild type                  | 24                 | 28                  | 0.537   |
| Mutant type                | 0                  | 1                   |         |
| Information absent         | 84                 | 97                  |         |
| BRAF                       |                    |                     |         |
| Wild type                  | 88                 | 87                  | 0.086   |
| Mutant type                | 1                  | 3                   |         |
| Information absent         | 19                 | 36                  |         |
| Metastatic site            |                    |                     |         |
| 1                          | 80                 | 39                  | <0.001* |
| >1                         | 28                 | 87                  |         |
| Diameter of metastasis, cm |                    |                     |         |
| <5                         | 94                 | 101                 | 0.159   |
| ≥5                         | 14                 | 25                  |         |

The symbol "\*" denotes P less than 0.05 in statistical comparison. LVI, lymphovascular invasion; PNI, perineural invasion.

Table 3 (continued)

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biopsy of primary tumor, are widely used for the prognosis evaluation of CRLM patients undergoing staged resection, but might not be appropriate for simultaneous resection (9,25-27).

According to the nomogram-based stratification, patients in low-risk group have a highly favorable outcome, and simultaneous resection is a rational option for such patients. Patients in high-risk group have a relatively poor outcome, therefore, further studies of the better surgical and therapeutic strategies for such patients are needed. There was a statistically significant difference in primary location, tumor grade, pT and pN stage, neural invasion, and metastatic site among the three groups. However, despite genetic predictors having been used with increasing frequency for patients with CRLM, the status of KRAS, NRAS, and BRAF did not differ obviously among risk groups (28,29).

The present study has several noteworthy limitations. First, our preoperative prognostic nomogram model was based on the retrospective data from a single clinical center, which may have biased the selection. Second, this model was internally validated with part of cases in our dataset but not externally validated. Third, incomplete genetic testing results and neoadjuvant therapy schemes might limit the accuracy of our conclusions. We do hope with the popularization of more accessible genetic testing method, molecular pathology will serve a prior part in future scoring systems. Further exploration is urgently needed to provide more precise risk assessment for CRLM patients.

#### Conclusions

This study formulated and validated a practicable preoperative prognostic nomogram model for surgeons and CRLM patients to predict individualized mortal risk after combined resection. We also designated cut-off to stratify patients at different level of risk to provide better assessment of postoperative survival. The CRLM patients ranked high-risk by the scoring system should consider comprehensive and more individualized treatment.

In conclusion, we identified five survival-associated preoperative factors of synchronous CRLM and established a prognostic model which can assess survival of CRLM patients undergoing simultaneous resection and exhibits high predictive power.

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#### Footnote

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#### References

- Bray F, Ferlay J, Soerjomataram I, et al. Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. CA Cancer J Clin 2018;68:394-424.
- Manfredi S, Lepage C, Hatem C, et al. Epidemiology and management of liver metastases from colorectal cancer. Ann Surg 2006;244:254-9.
- 3. Glimelius B, Påhlman L, Cervantes A; ESMO Guidelines Working Group. Rectal cancer: ESMO Clinical Practice

Guidelines for diagnosis, treatment and follow-up. Ann Oncol 2010;21 Suppl 5:v82-6.

- 4. Zhu D, Ren L, Wei Y, et al. Outcome of patients with colorectal liver metastasis: analysis of 1,613 consecutive cases. Ann Surg Oncol 2012;19:2860-8.
- Baltatzis M, Chan AK, Jegatheeswaran S, et al. Colorectal cancer with synchronous hepatic metastases: Systematic review of reports comparing synchronous surgery with sequential bowel-first or liver-first approaches. Eur J Surg Oncol 2016;42:159-65.
- Kelly ME, Spolverato G, Lê GN, et al. Synchronous colorectal liver metastasis: a network meta-analysis review comparing classical, combined, and liver-first surgical strategies. J Surg Oncol 2015;111:341-51.
- Idrees JJ, Bagante F, Gani F, et al. Population level outcomes and costs of single stage colon and liver resection versus conventional two-stage approach for the resection of metastatic colorectal cancer. HPB (Oxford) 2019;21:456-64.
- Ghiasloo M, Pavlenko D, Verhaeghe M, et al. Surgical treatment of stage IV colorectal cancer with synchronous liver metastases: A systematic review and network metaanalysis. Eur J Surg Oncol 2020;46:1203-13.
- Fong Y, Fortner J, Sun RL, et al. Clinical score for predicting recurrence after hepatic resection for metastatic colorectal cancer: analysis of 1001 consecutive cases. Ann Surg 1999;230:309-18; discussion 318-21.
- Margonis GA, Sasaki K, Gholami S, et al. Genetic And Morphological Evaluation (GAME) score for patients with colorectal liver metastases. Br J Surg 2018;105:1210-20.
- Collins GS, Reitsma JB, Altman DG, et al. Transparent reporting of a multivariable prediction model for individual prognosis or diagnosis (TRIPOD): the TRIPOD Statement. BMC Med 2015;13:1.
- Edge SB, Compton CC. The American Joint Committee on Cancer: the 7th edition of the AJCC cancer staging manual and the future of TNM. Ann Surg Oncol 2010;17:1471-4.
- Church R, Fleshman J, Dietz D, et al. Liver resection for colorectal metastases. Ann R Coll Surg Engl 2004;86:401; author reply 401-2.
- 14. Thelen A, Jonas S, Benckert C, et al. Simultaneous versus staged liver resection of synchronous liver metastases from colorectal cancer. Int J Colorectal Dis 2007;22:1269-76.
- Hillingsø JG, Wille-Jørgensen P. Staged or simultaneous resection of synchronous liver metastases from colorectal cancer--a systematic review. Colorectal Dis 2009;11:3-10.
- 16. Sasaki K, Margonis GA, Andreatos N, et al. Pre-

hepatectomy carcinoembryonic antigen (CEA) levels among patients undergoing resection of colorectal liver metastases: do CEA levels still have prognostic implications? HPB (Oxford) 2016;18:1000-9.

- Ueno H, Konishi T, Ishikawa Y, et al. Prognostic value of poorly differentiated clusters in the primary tumor in patients undergoing hepatectomy for colorectal liver metastasis. Surgery 2015;157:899-908.
- Liu W, Wang HW, Wang K, et al. The primary tumor location impacts survival outcome of colorectal liver metastases after hepatic resection: A systematic review and meta-analysis. Eur J Surg Oncol 2019;45:1349-56.
- Al Bandar MH, Kim NK. Current status and future perspectives on treatment of liver metastasis in colorectal cancer (Review). Oncol Rep 2017;37:2553-64.
- 20. Nordlinger B, Van Cutsem E, Rougier P, et al. Does chemotherapy prior to liver resection increase the potential for cure in patients with metastatic colorectal cancer? A report from the European Colorectal Metastases Treatment Group. Eur J Cancer 2007;43:2037-45.
- 21. Viganò L, Capussotti L, De Rosa G, et al. Liver resection for colorectal metastases after chemotherapy: impact of chemotherapy-related liver injuries, pathological tumor response, and micrometastases on long-term survival. Ann Surg 2013;258:731-40; discussion 41-2.
- 22. Khan AZ, Morris-Stiff G, Makuuchi M. Patterns of chemotherapy-induced hepatic injury and their implications for patients undergoing liver resection for colorectal liver metastases. J Hepatobiliary Pancreat Surg 2009;16:137-44.
- Zheng P, Chen Q, Li J, et al. Prognostic Significance of Tumor Deposits in Patients With Stage III Colon Cancer: A Nomogram Study. J Surg Res 2020;245:475-82.
- 24. Shang-Guan XC, Chen QY, Li P, et al. Preoperative lymph node size is helpful to predict the prognosis of patients with stage III gastric cancer after radical resection. Surg Oncol 2018;27:54-60.
- 25. Rees M, Tekkis PP, Welsh FK, et al. Evaluation of longterm survival after hepatic resection for metastatic colorectal cancer: a multifactorial model of 929 patients. Ann Surg 2008;247:125-35.
- 26. Nordlinger B, Guiguet M, Vaillant JC, et al. Surgical resection of colorectal carcinoma metastases to the liver. A prognostic scoring system to improve case selection, based on 1568 patients. Association Française de Chirurgie. Cancer 1996;77:1254-62.
- 27. Iwatsuki S, Dvorchik I, Madariaga JR, et al. Hepatic resection for metastatic colorectal adenocarcinoma: a

proposal of a prognostic scoring system. J Am Coll Surg 1999;189:291-9.

28. Søreide K, Sandvik OM, Søreide JA. KRAS mutation in patients undergoing hepatic resection for colorectal liver metastasis: a biomarker of cancer biology or a byproduct

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29. De Cuyper A, Van Den Eynde M, Machiels JP. HER2 as a Predictive Biomarker and Treatment Target in Colorectal Cancer. Clin Colorectal Cancer 2020;19:65-72.



Figure S1 Kaplan-Meier curve showing stable procedure outcome through follow-up period.